

**LEUKOTRIENE RECEPTOR ANTAGONISTS VERSUS
PLACEBO IN ADULTS AND ADOLESCENTS WITH ASTHMA:
A SYSTEMATIC REVIEW AND META-ANALYSIS**

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Abstract

Background: Leukotriene receptor antagonists (LTRAs) are recommended as alternative treatment in patients with mild asthma, but their relative effect compared with placebo is unknown. **Objective:** To determine the benefits and harms of LTRAs compared with placebo in adults and adolescents with asthma. **Data sources:** MEDLINE and the Cochrane Central Register of Controlled Trials from inception through December 2012. **Study selection:** Peer-reviewed, English-language, randomized controlled trials (≥ 4 weeks in duration) in patients with asthma that reported the effect of LTRAs versus placebo on measures of asthma control. **Data extraction:** Data on the study population, interventions, outcome measures, adverse events, and study methodology were extracted by three authors. **Data synthesis:** Forty seven trials satisfied our eligibility criteria. Random-effects model meta-analyses, random-effects meta-regression, and subgroup analyses were performed. In 9 trials, LTRAs reduced the risk of an exacerbation by 35% (summary risk ratio = 0.65, 95% CI 0.50, 0.84). The effect was more pronounced in studies of shorter duration ($p < 0.01$). LTRAs significantly increased FEV₁ (summary mean difference [MD] from 13 trials = 0.11 liters, 95% CI 0.08, 0.15; summary MD in percent change from 11 trials = 5.95, 95% CI 3.3, 8.6) and FEV₁ % predicted (summary MD from 8 trials = 4.2%, 95% CI 1.5, 1.9). Daytime symptoms (summary standardized MD from 14 trials = -0.21, 95% CI -0.37, -0.04), short-acting β_2 -agonist use (summary MD from 11 trials = -0.65 puffs/day, 95% CI -0.82, -0.49; summary MD in percent change from 8 trials = -16.4, 95% CI -22.4, -10.4), nocturnal awakenings (summary MD from 7 trials = -0.66, 95% CI -1, -0.3), and asthma-specific quality of life (summary standardized MD from 5 trials = 0.13, 95% CI 0.02, 0.2) were also significantly improved compared to placebo. The

proportions of patients with adverse events were similar between intervention and comparator groups. **Limitations:** Variation in definitions and reporting of outcomes, large heterogeneity, and possible selective outcome reporting bias. **Conclusion:** LTRAs improved asthma control compared to placebo. It remains unclear however, which patients with asthma are more likely to respond to treatment with LTRAs.

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List of abbreviations

AE: Adverse Events
AIA: Aspirin Induced Asthma
ALT: Alanine Aminotransferase
AQLQ: Asthma Quality of Life Questionnaire
AR: Allergic Rhinitis
ASQL: Asthma Specific Quality of Life
AST: Aspartate Aminotransferase
BEC: Beclomethasone
BUD: Budesonide
CI: Confidence Interval
ER: Emergency Room
FEV₁: Forced Expiratory Volume in one second
FCS: Fluticasone
ICS: Inhaled Corticosteroids
ITT: Intention To Treat analysis
LABA: Long Acting β_2 Agonist
LTRA: Leukotriene Receptor Antagonist
MD: Mean Difference
PEFR: Peak Expiratory Flow Rate
RCT: Randomized Controlled Trial
RR: Risk Ratio
SABA: Short Acting β_2 Agonist
SD: Standard Deviation
SMD: Standardized Mean Difference
SRDR: Systematic Review Data Repository
URTI: Upper Respiratory Tract Infection

Introduction

Asthma is one of the most common chronic diseases with considerable social and economic burden, involving both high direct costs related to healthcare utilization and indirect costs related to time lost from work or school. In the US the annual cost is estimated around \$18 billion.¹ Approximately 300 million people worldwide, and 25 million Americans, are affected by asthma. These numbers are expected to rise to 400 million by 2025.²

The successful long-term management of asthma includes the use of medications that target the underlying inflammatory process. Although inhaled corticosteroids (ICS) constitute the current gold standard of maintenance treatment, leukotriene receptor antagonists (LTRAs) have the advantage of being administered orally in a single or twice daily dose; importantly, these agents appear to lack the adverse effects associated with long-term corticosteroid therapy.³ In addition, their mechanism of action theoretically predicts a good response in patients with specific asthma ‘phenotypes’. Allergic rhinitis (AR) is present in many patients with asthma and LTRAs might improve asthma-related outcomes by treating both conditions concurrently.⁴ Moreover, aspirin-induced asthma (AIA), which is clinically characterized by chronic eosinophilic rhinosinusitis, nasal polyposis, aspirin hypersensitivity, and development of persistent asthma, is associated with increased airway leukotrienes and is frequently poorly responsive to ICS.⁵ Current guidelines recommend the use of LTRAs as monotherapy in patients with mild persistent asthma, as an alternative, or as add-on therapy to ICS, and as an alternative to either increasing the ICS dose or adding a long-acting β_2 -agonist.⁶ However, the relative benefits and harms of LTRAs compared with placebo have not been established.

We conducted a systematic review of randomized controlled trials (RCTs) that compared the efficacy and safety of LTRAs with placebo in adults and adolescents with asthma for both objective and patient-reported outcome measures used to assess asthma control.

Materials and Methods

Data sources and search

We searched MEDLINE and the Cochrane Central Register of Controlled trials from inception through December 2012. We developed a search strategy with a combination of Medical Subject Headings terms and free text keywords relevant to study design (“randomized controlled trial”), disease of interest (“asthma”), and intervention of interest (“leukotriene receptor antagonists”) [Table 1].

Eligibility criteria and study selection

We included peer-reviewed publications of RCTs if they fulfilled the following criteria: comparison of a LTRA either as monotherapy or as add-on therapy to ICS with placebo in adults and adolescents (≥ 12 years) with asthma; oral administration of usual licensed doses of a LTRA on a daily basis (montelukast 10 mg once daily for individuals 15 years and older, zafirlukast 20 mg twice daily for individuals 12 years and older, pranlukast 225 mg twice daily for individuals 12 years and older); minimum treatment duration of 4 weeks; inclusion of at least one pre-specified outcome measure that reflects asthma control (asthma exacerbations, pulmonary function tests, daytime asthma symptom scores, asthma-specific quality of life, nocturnal awakenings, short acting β_2 -agonist use, adverse events); and English language publication. The primary outcome measure was the number of

exacerbations requiring systemic corticosteroids, an unscheduled visit to a doctor, or a visit to an emergency department. Asthma-specific quality of life is assessed using the asthma-specific quality of life and mini asthma quality of life questionnaires.⁷⁻⁸ The scales range from 1 to 7 (or 0 to 6), with higher values indicating better quality of life. The minimally important difference considered clinically important is 0.5.⁷⁻⁸ Due to the inclusion of children and adolescents in some studies of montelukast, we included studies in which at least some children and adolescents received 10 mg daily and excluded those in which none of the participants received 10 mg. Two investigators independently reviewed the titles and abstracts of the citations for potentially relevant articles using Abstrackr;⁹ the full text publications of potentially relevant articles were retrieved and rescreened by the same two investigators. Disagreements were resolved by consensus.

Data extraction

Each eligible study was independently data extracted by two of three investigators; any disagreements were resolved by consensus. We extracted data on study design and methodology, patient characteristics, interventions, comparators, outcome measures, and adverse events using a standardized electronic form in the Systematic Review Data Repository (SRDR), which is an open-access, collaborative, Web-based repository of systematic review data.¹⁰

Assessment of risk of bias

We assessed the methodological quality of the eligible studies using the Cochrane Collaboration's 'Risk of bias' tool.¹¹ This tool includes 13 'risk of bias' items (Table 2). A judgment of 'low', 'high', or 'unclear' risk of bias was assigned for the first seven items (sequence generation, allocation concealment, patients' blinding, caregivers' blinding,

outcome assessors' blinding, attrition, selective outcome reporting), whereas a judgment of 'yes', 'no', or 'unsure' was assigned for the remainder (intention-to-treat analysis, baseline balance, co-interventions similarity, compliance, presence of other biases). Reviewing across all risk of bias items, we assigned an overall quality grade of good, fair, or poor to each RCT. We considered a study of poor quality if any of the following was observed: a) absence of blinding, b) differential loss-to-follow up, c) baseline imbalances, d) absence of a washout period in the case of crossover trials. Studies that reported sufficient details about the implementation of blinding (e.g., double-blind and use of identical capsules) were considered as having low risk of bias for this specific item, whereas studies with insufficient reporting (e.g., double-blind) were considered as having unclear risk of bias. Blinding of the outcome assessors with regard to patient-reported outcomes was considered adequate if patients were reported to be blinded. We compared the proportions of withdrawals in each group using the Chi-square test and a p-value less than 0.1 was indicative of differential loss-to-follow up. An analysis was considered as intention-to-treat (ITT) if the number of participants who were randomized was equal to the number of participants who were analyzed. The assessments were completed by one author.

Data synthesis and analysis

Study treatment effects for binary outcomes were estimated using the risk ratio (RR) with corresponding 95% confidence intervals (CI). For the continuous outcome measures, the difference in mean changes from baseline between LTRAs and placebo was calculated for each study. The 95% CI was calculated based on the pooled standard deviation (SD) of calculated differences. Study treatment effects for daytime symptom scores and asthma-specific quality of life scores were computed using Hedges' g statistic corrected for small

samples.¹² When only the baseline and final SDs were reported, we calculated SDs of change from baseline in each group assuming a correlation coefficient of 0.5.¹³ When means and measures of dispersion were not reported in the text, they were approximated from figures using Engauge Digitizer Qt4.¹⁴ We imputed missing group SDs in one study using the median of all available SDs from other studies.

In the meta-analyses, the summary treatment effects were estimated using the random-effects model estimated by restricted maximum likelihood.¹⁵ Random-effects modeling assumes a genuine diversity in the results of various studies and incorporates a between-study variance in the calculations. We calculated a summary RR, a summary mean difference, and a summary standardized mean difference (SMD) between LTRAs and placebo, where appropriate. Statistical heterogeneity was quantified by the I^2 statistic.¹⁶ Values around 25%, 50%, and 75% indicate low, moderate, and high heterogeneity, respectively.¹⁶ The overall analysis included only the reported endpoint values at the longest follow-up within each trial. Subgroup analyses and random effects meta-regression were employed to explore the effect of pre-specified factors on the effect estimates, when an outcome of interest was reported by at least three RCTs in each subgroup. The pre-specified factors were: type of LTRA, use of ICS, dose of ICS, treatment duration, asthma severity, presence of comorbid allergic rhinitis, aspirin-induced asthma. The RCTs were classified into three categories based on the concomitant use of ICS in the intervention groups; no use of ICS, equal use of ICS, or unequal use of ICS. Wherever possible, doses of ICS were converted to microfine hydrofluoroalkane-beclomethasone dipropionate (HFA-BDP) equivalent based on 1 µg of microfine HFA-propelled beclomethasone = 2 µg of chlorofluorocarbon (CFC) – propelled beclomethasone = 1 µg of fluticasone = 2 µg of

budesonide = 1 µg of ciclesonide = 1 µg of mometasone = 4 µg of triamcinolone = 4 µg of flunisolide.¹⁷ Since the treatment duration varied among the trials, we grouped the time points of outcome assessments in the individual trials into six intervals: 4-7, 8-11, 12-15, 16-23, 24-30, and more than 30 weeks. This grouping was designed to best capture all available data in the trials, but also reflects periodic monitoring of asthma control used in clinical practice. Crossover trials were not included in the primary meta-analyses. All analyses were performed with OpenMetaAnalyst.¹⁸

Sensitivity analyses

Additional analyses addressed: 1) trials in which all participants were 12 years and older in order to investigate the impact of including RCTs with overlapping populations of children and adolescents; 2) trials with an unclear definition of exacerbation; and 3) inclusion of two crossover trials that assessed the outcomes FEV₁, short-acting β₂-agonist use, and nocturnal awakenings in paired analyses.

Table 1. Search strategy

1. Leukotriene antagonists.sh.
 2. Leukotriene receptor antagonist*.af.
 3. (leukotriene and receptor antagonist*.af.
 4. Leukotriene modifier*.af.
 5. (leukotriene receptor and antagonist*).af.
 6. Montelukast.af.
 7. Zafirlukast.af.
 8. Pranlukast.af.
 9. Singulair.af.
 10. Accolate.af.
 11. Onon.af.
 12. Azlaire.af.
 13. Or/1-12
 14. Exp Asthma/
 15. Asthma*.af.
 16. Samter* syndrome.af.
 17. Aspirin intolerance.af.
 18. Aspirin sensitivity.af.
 19. Aspirin hypersensitivity.af.
 20. Exercise induced broncho*.af.
 21. (Exercise induced and broncho*).af.
 22. Nasal polyp*.af.
 23. Or/14-22
 24. Randomized controlled trial.pt.
 25. Controlled clinical trial.pt.
 26. Randomized controlled trials/
 27. Random Allocation/
 28. Double-blind method/
 29. Single-blind method/
 30. Clinical trial.pt.
 31. Clinical Trials.mp. or exp Clinical Trials/
 32. (clinic\$ adj25 trial\$.tw.
 33. ((singl\$ or doubl\$ or treb\$ or tripl\$) adj (mask\$ or blind\$)).tw.
 34. Placebos/
 35. Placebo\$.tw.
 36. Random\$.tw.
 37. Trial\$.tw.
 38. (randomized control trial or clinical control trial).sd.
 39. Latin adj square.tw.
 40. Comparative Study.tw. or Comparative Study.pt.
 41. Exp Evaluation studies/
 42. Follow-up Studies/
 43. Prospective Studies/
 44. (control\$ or prospective\$ or volunteer\$.tw.
 45. Cross-over Studies/
 46. Or/24-45
 47. And/13, 23, 46
-

Table 2. Risk of bias items assessed for randomized controlled trials

1. What is the risk of selection bias (biased allocation to interventions) due to inadequate generation of a randomized sequence? [Low, Unclear, High]
2. What is the risk of selection bias (biased allocation of interventions) due to inadequate concealment of allocations before assignment? [Low, Unclear, High]
3. For each main outcome or class of outcomes, what is the risk of performance bias due to knowledge of the allocated interventions by participants and personnel during the study (lack of study participant and personnel blinding)? [Low, Unclear, High]
4. Was the care provider blinded to the intervention? [Low, Unclear, High]
5. For each main outcome or class of outcomes, what is the risk of detection bias due to knowledge of the allocated interventions by outcome assessment (lack of outcome assessor blinding)? [Low, Unclear, High]
6. For each main outcome or class of outcomes, what is the risk of attrition bias due to amount, nature, or handling of incomplete outcome data? [Low, Unclear, High]
7. What is the risk of reporting bias due to selective outcome reporting? [Low, Unclear, High]
8. Were all randomized participants analyzed in the group to which they were allocated? [Yes, No, Unsure].
9. Were the groups similar at baseline regarding the most important prognostic indicators? [Yes, No, Unsure]
10. Were co-interventions avoided or similar? [Yes, No, Unsure]
11. Was the compliance acceptable in all groups? [Yes, No, Unsure]
12. Was the timing of the outcome assessment similar in all groups?*[Yes, No, Unsure]
13. Are there other risks of bias? [Yes, No]

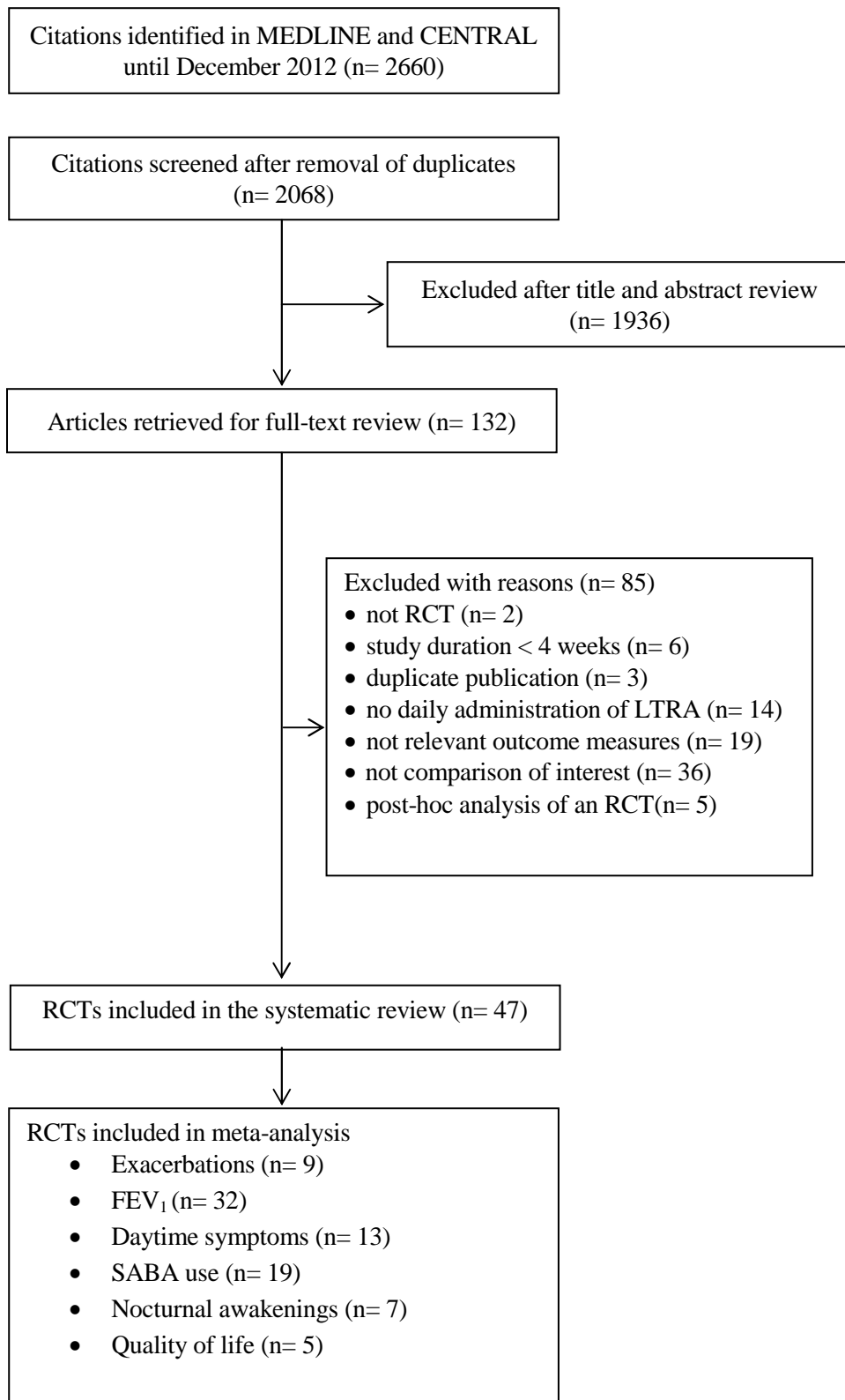
* Question 12 was incorporated into question 13 for the purpose of the current study

Results

Literature search

Figure 1 summarizes our search yield. We screened 2068 citations. A total of 224 articles were retrieved for full-text review, and 47 RCTs¹⁹⁻⁶⁵ met our inclusion criteria.

Figure 1. Flow diagram



Trial characteristics

Table 3 summarizes the characteristics of the included RCTs, which were published between 1994 and 2011. Five RCTs included both children and adolescents.^{51,54-56,58}

Asthma was generally diagnosed by the demonstration of reversibility of airway obstruction after administration of short-acting β_2 -agonists. The main exclusion criteria within trials were active upper respiratory tract infection, recent use of oral corticosteroids, and recent ER visit or hospitalization due to worsening asthma. Smokers were excluded in 28 trials.

Overall, 9057 patients were randomized to receive either a LTRA or placebo. Montelukast was administered in 33 RCTs, zafirlukast in 9, and pranlukast in 5. ICS were used as concomitant treatments by all patients in 15 trials, whereas in 8 trials ICS were used only by a proportion of participants. In 4 RCTs the dose of ICS was gradually reduced during follow-up according to specific criteria described in the trials. Short-acting β_2 -agonists were permitted on an “as needed” basis in every trial. Mean FEV₁ at baseline was between 59% and 102% of predicted values. Sixteen RCTs reported inclusion of patients with a history of atopy, 9 reported presence of concomitant allergic rhinitis, and 2 RCTs reported inclusion of patients with aspirin-induced asthma. Patients with exercise induced bronchoconstriction were included in 5 trials. There were 38 parallel and 9 crossover RCTs. Twenty eight RCTs were multicenter. Treatment duration ranged from 4 to 30 weeks.

Table 3. Characteristics of included trials

| Source | Region | Treatment arms | Concomitant treatments | Treatment duration, weeks | Patients, n | Age, years | Sex, %Male | FEV ₁ % predicted at baseline (SD) | FEV ₁ % predicted range of inclusion |
|--|-----------------------|---|-------------------------------------|---------------------------|-------------|------------|------------|---|---|
| Altman 1998 ¹⁹ | USA | Montelukast | Theophylline, ICS*, | 6 | 57 | 33 median | 79 | 62 (13) | 40-80 |
| | | Placebo | SABA | | 58 | 36 median | 78 | 59 (13) | |
| American Lung Association 2007 ²⁰ | USA | Montelukast | ICS*, LABA, SABA | 24 | 164 | 40 | 28 | 77 (17) | > 50 |
| | | Placebo | | | 164 | 40 | 26 | 80 (16) | |
| Awad 2002 ²¹ | India | Zafirlukast | SABA | 12 | 116 | 35 | 47 | 62 (11) | 45-80 |
| | | Placebo | | | 99 | 35 | 50 | 64 (12) | |
| Baena-Cagnani 2003 ²² | USA | Montelukast | SABA | 4 | 311 | 34 | 39 | 86 | > 70 |
| | | Placebo | | | 302 | 32 | 33 | 86 | |
| Barnes 1997 ²³ | Europe | Pranlukast | BEC* ($\leq 1000 \mu\text{g/d}$), | 4 | 46 | 39 | 59 | 68 (12) | 50-80 |
| | | Placebo | SABA | | 44 | 38 | 66 | 67 (11) | |
| Baumgartner 2003 ²⁴ | North & South America | Montelukast | SABA | 6 | 313 | 36 | 34 overall | 69 (12) | 50-85 |
| | | Placebo | | | 103 | 36 | | 68 (12) | |
| Busse 2001 ²⁵ | USA | Zafirlukast | SABA | 12 | 111 | 12-75 | 50 overall | 66-67 overall | 50-80 |
| | | Placebo | | | 114 | overall | | | |
| Cakmak 2004 ²⁶ | Turkey | Zafirlukast/ BUD (400 $\mu\text{g/d}$) | SABA | 6 | 11 | 30 | 55 | 88 (14) | ≥ 70 |
| | | Placebo/ BUD (400 $\mu\text{g/d}$) | | | 10 | 28 | 20 | 89 (14) | |
| Dahlén 2002 ²⁷ | USA, Europe | Montelukast | Theophylline, ICS*, | 4 | 40 40 | 49 median | 38 | 70 | ND |
| | | Placebo | SABA | | | 47 median | 28 | 70 | |
| Fish 1997 ²⁸ | USA | Zafirlukast | SABA | 13 | 514 | 18-55 | 57 | 78 (16) | ≥ 55 |
| | | Placebo | | | 248 | (80%) | 59 | 79 (17) | |
| Green 2006 ²⁹ | UK | Montelukas/BUD (200 $\mu\text{g/d}$) | SABA | 4 | 49 | 42 median | 51 | 75 (3) | ND |
| | | Placebo / BUD (200 $\mu\text{g/d}$) | | | | | | | |
| Helenius 2004 ³⁰ | Finland | Montelukast | SABA | 4 | 16 | 18 | 100 | 101 (12) | ND |
| | | Placebo | | | | | | | |
| Huang 2003 ³¹ | Taiwan | Zafirlukast | BUD (800-1600 $\mu\text{g/d}$), | 4 | 20 | 59 | 53 | 68 (1) | 60-80 |
| | | Placebo | SABA | | 18 | 57 | 50 | 69 (1) | |
| Israel 2002 ³² | USA | Montelukast | SABA, antihistamines | 6 | 339 | 34 | 48 | 67 (11) | 50-80 |
| | | Placebo | | | 111 | 33 | 47 | 67 (12) | |
| Jayaram 2005 ³³ | Canada | Montelukast | BUD (1857 $\mu\text{g/d}$), | 4 | 14 | 61 | 57 | 62 (15) | ND |
| | | Placebo | SABA | | | | | | |
| Jayaram 2005 ³⁴ | Canada, Brazil | Montelukast | SABA | 8 | 19 | 31 | 58 | 77 (16) | ND |
| | | Placebo | | | 13 | 39 | 71 | 80 (23) | |
| Kanazawa 2004 ³⁵ | Japan | Pranlukast/ BEC (800 $\mu\text{g/d}$) | SABA | 4 | 10 | 28 | 60 | 87 | ND |
| | | Placebo/ BEC (800 $\mu\text{g/d}$) | | | | | | | |
| Kanniess 2002 ³⁶ | Germany | Montelukast | BEC (tapered doses), | 12 | 26 | 38 | 50 | 95 (10) | > 80 |
| | | Placebo | SABA | | 24 | 43 | 46 | 92.3 (9) | |
| Kraft 2006 ³⁷ | USA | Montelukast | SABA | 4 | 19 | 38 | 32 | 83 (3) | ND |
| | | Placebo | | | | | | | |

Table 3. (continued)

| Source | Region | Treatment arms | Concomitant treatments | Treatment duration, weeks | Patients, n | Age, years | Sex, %Male | FEV ₁ % predicted at baseline (SD) | FEV ₁ % predicted range of inclusion |
|--------------------------------------|--|------------------------------|---------------------------|---------------------------|-------------|------------|------------|---|---|
| Laviolette 1999 ³⁸ | North America, Europe, Africa, Australia, Asia | Montelukast/ BEC (400 µg/d) | SABA | 16 | 193 | 40 median | 56 | 72 (12) | 50-85 |
| | | Placebo/ BEC (400 µg/d) | | | 200 | 39 median | 52 | 71 (12) | |
| | | Montelukast | | | 201 | 38 | 49 | 72 (12) | |
| | | Placebo | | | 48 | 41 | 40 | 71 (11) | |
| Leff 1998 ³⁹ | USA | Montelukast | SABA, | 12 | 54 | 25 | 52 | 83 (11) | ND |
| | | Placebo | antihistamines | | 56 | 25 | 52 | 84 (11) | |
| Löfdahl 1999 ⁴⁰ | USA, Canada, Europe | Montelukast | ICS (various doses), SABA | 12 | 113 | 40 | 42 | 85 (11) | > 70 |
| | | Placebo | | | 113 | 41 | 54 | 82 (13) | |
| Malmstrom 1999 ⁴¹ | Europe, Africa, Australia, Central and South America | Montelukast | theophylline, | 12 | 387 | 35 median | 40 | 65 (10) | 50-85 |
| | | Placebo | SABA | | 257 | 36 median | 43 | 66 (11) | |
| Minoguchi 2002 ⁴² | Japan | Montelukast | Theophylline, | 4 | 26 | 37 | 50 | 83 (16) | ND |
| | | Placebo | SABA | | | | | | |
| Nakamura 1998 ⁴³ | Japan | Pranlukast | SABA | 4 | 10 | 35 median | 80 | 71 median | ≥ 50 |
| | | Placebo | | | 7 | 32 median | 43 | 80 median | |
| Nathan 1998 ⁴⁴ | USA | Zafirlukast | SABA, nasal | 13 | 231 | 33 | 45 | 66.6 overall | 45-80 |
| | | Placebo | corticosteroids | | 223 | 32 | 41 | | |
| Nathan 1999 ⁴⁵ | USA | Zafirlukast | SABA | 13 | 96 | 32 | 55 | 77 (15) | > 55 |
| | | Placebo | | | 95 | 30 | 53 | 78 (17) | |
| Nathan 2005 ⁴⁶ | USA | Montelukast | FCS/LABA | 4 | 282 | 34 | 33 | 81 (10) | ND |
| | | Placebo | (100/50 µg/d), SABA | | 290 | 36 | 28 | 81 (10) | |
| Pizzichini 1999 ⁴⁷ | USA, Canada | Montelukast | SABA | 4 | 19 | 31 | 63 | 69 (11) | ND |
| | | Placebo | | | 21 | 28 | 57 | 69 (15) | |
| Reid (A) 2008 ⁴⁸ | Australia | Zafirlukast | SABA | 12 | 14 | 42 median | 57 | 85 median | ≥ 60 |
| | | Placebo | | | 7 | 29 median | 43 | 80 | |
| Reid (B) | | Zafirlukast/ BUD (1600 µg/d) | | | 16 | 37 median | 56 | 77 | |
| | | Placebo/ BUD (1600 µg/d) | | | 8 | 45 median | 25 | 76 | |
| Reiss 1998 ⁴⁹ | USA | Montelukast | ICS*, SABA | 12 | 408 | 31 median | 43 | 67 (11) | 50-85 |
| | | Placebo | | | 273 | 31 median | 47 | 69 (11) | |
| Schäper 2011 ⁵⁰ | Germany | Montelukast | ICS*, SABA | 6 | 24 | 56 median | 71 | 88 | ND |
| | | Placebo | | | | | | | |
| Spahn 2006 ⁵¹ | USA | Montelukast | SABA | 8 | 11 | 13 | 64 | 88 (10) | 60-90 |
| | | Placebo | | | 10 | 14 | 36 | 83 (10) | |

Table 3. (continued)

| Source | Region | Treatment arms | Concomitant treatments | Treatment duration, weeks | Patients, n | Age, years | Sex, %Male | FEV ₁ % predicted at baseline (SD) | FEV ₁ % predicted range of inclusion |
|--|-------------|-----------------------------------|--------------------------------|---------------------------|-------------|-------------|------------|---|---|
| Spector 1994 ⁵² | USA | Zafirlukast | SABA | 6 | 70 | 37 | 74 | 66 | 40-75 |
| | | Placebo | | | 70 | 36 | 71 | 69 | |
| Spector 2004 ⁵³ | USA | Montelukast | SABA, nasal corticosteroids | 4 | 8 | 42 | 50 | > 50 | 50-85 |
| | | Placebo | | | 6 | 36 | 67 | | |
| Stelmach (A) 2007 ⁵⁴ | Poland | Montelukast/ BUD (200µg/d) | SABA | 4 | 29 | 11 | 69 | 95 (11) | ND |
| | | Placebo/ BUD (200µg/d) | | | 29 | 12 | 69 | 94 (10) | |
| Stelmach (B) | | Montelukast | | | 29 | 10 | 62 | 96 (11) | |
| | | Placebo | | | 29 | 11 | 70 | 95 (10) | |
| Stelmach 2002 ⁵⁵ | Poland | Montelukast | SABA | 4 | 18 | 12 | 60 | 77 (4) | ND |
| | | Placebo | | | 36 | 12 | 44 | 75 (5) | |
| Stelmach 2002 ⁵⁶ | Poland | Montelukast | SABA | 6 | 16 | 14 | 67 | 85.(9) | ND |
| | | Placebo | | | 19 | 13 | 59 | 81 (7) | |
| Storms 2004 ⁵⁷ | USA | Montelukast/ FCS (200µg/d) | SABA | 4 | 39 | 33 | 28 | 88 (10) | ≥ 70 |
| | | Placebo/ FCS (200µg/d) | | | 44 | 31 | 45 | 88 (11) | |
| Strunk 2008 ⁵⁸ | USA | Montelukast | BUD (800-1600µg/d), LABA, SABA | 30 | 19 | 11 | 58 | 102 (14) | ND |
| | | Placebo | | | 19 | | | | |
| Tohda 2002 ⁵⁹ | Japan | Montelukast | BEC (various doses), SABA | 24 | 93 | 16-70 range | 58 | 87 (18) | ND |
| | | Placebo | | | 98 | | 58 | 86 (25) | |
| Ulrik 2009 ⁶⁰ | Denmark | Montelukast | SABA | 12 | 16 | 34 | 44 | 79 (14) | > 70 |
| | | Placebo | | | 15 | 33 | 40 | 83 (10) | |
| Vaquerizo 2003 ⁶² | Spain | Montelukast / BUD (400 -1600µg/d) | SABA | 16 | 326 | 42 | 62 | 81 (19) | ≥ 55 |
| | | Placebo/ BUD (400 -1600µg/d) | | | 313 | 44 | 61 | 81 (21) | |
| Wise (A) 2009 ⁶³ | USA | Montelukast | ICS*, SABA | 4 | 120 | 37 | 19 | 87 (12) | > 75 |
| | | Placebo | | | 121 | 39 | 27 | 87 (15) | |
| Wise (B) | | Montelukast | | | 119 | 37 | 29 | 86 (13) | |
| | | Placebo | | | 120 | 39 | 32 | 87 (13) | |
| Yoo 2001 ⁶³ | South Korea | Pranlukast | ICS*, SABA | 4 | 98 | 45 | 61 | 73 | 60-80 |
| | | Placebo | | | 99 | 45 | 54 | 73 | |
| Yoshida 2002 ⁶⁴ | Japan | Pranlukast | Theophylline, SABA | 4 | 32 | 41 | 44 | 77 (7) | ND |
| | | Placebo | | | | | | | |
| Zeidler 2006 ⁶⁵ | USA | Montelukast | SABA | 4 | 20 | 36 | 56 | 86 (12) | ≥ 60 |
| | | Placebo | | | | | | | |

Abbreviations: BEC: Beclomethasone; BUD: Budesonide; FCS: Fluticasone; ICS: Inhaled corticosteroids; ND: No data; SABA: Short acting β₂-agonists; SD: Standard deviation.

* ICS used only by a proportion of participants (20%-90%)^a Data are shown as means, unless otherwise specified.

Assessment of risk of bias

Among the 47 RCTs, 20 were assessed to be of good quality, 13 of fair quality, and 14 of poor quality (Table 4). Generation of a randomized sequence and allocation concealment were not clearly reported in the majority of trials. All but one RCT reported double-blinding; the other RCT was single-blind. Six RCTs had differential loss-to-follow-up. Intention-to-treat analyses were not widely used. RCTs were generally balanced with regard to baseline characteristics and co-interventions. Among the nine crossover trials, six had at least a 1-week wash-out period which was considered adequate; one had no wash-out period and two did not clarify whether a wash-out period was implemented.

Table 5 comprises all included trials and shows which studies contributed data to the meta-analysis of each outcome.

Table 4. Risk of bias in included trials

| Source | Sequence generation | Allocation concealment | Blinding/ Patients | Blinding/ Caregivers | Blinding/ Assessors | Attrition | Selective outcome reporting | Intention-to-treat analysis | Baseline balance | Co-interventions similarity | Compliance | Other biases | Overall grade |
|---------------------------------------|---------------------|------------------------|--------------------|----------------------|---------------------|-----------|-----------------------------|-----------------------------|------------------|-----------------------------|------------|--------------------|---------------|
| Altman 1998 | Unclear | Unclear | Low | Low | Low | Unclear | Low | Yes | Yes | Yes | Unclear | No | Fair |
| American Lung Association 2007 | Low | Unclear | Low | Low | Low | Low | Unclear | No | Yes | Yes | Yes | No | Good |
| Awad 2002 | Unclear | Unclear | Unclear | Unclear | Unclear | Low | Unclear | Yes | Yes | Unclear | Unclear | No | Fair |
| Baegna-Cagnani 2003 | Low | Unclear | Low | Low | Low | High | Unclear | Yes | Yes | Yes | Unclear | No | Poor |
| Barnes 1997 | Unclear | Unclear | Low | Low | Low | Low | Low | Yes | Yes | Yes | Unclear | No | Good |
| Baumgartner 2003 | Low | Unclear | Low | Low | Low | Low | Unclear | Yes | Yes | Yes | Yes | No | Good |
| Busse 2001 | Low | Unclear | Low | Low | Low | Low | Unclear | Yes | Yes | Yes | Unclear | No | Good |
| Cakmak 2004 | High | Unclear | Unclear | Unclear | Unclear | Unclear | High | Unclear | No | Unclear | Unclear | No | Poor |
| Dahlén 2002 | Low | Unclear | Low | Low | Low | Low | Unclear | Yes | Yes | Yes | Unclear | No | Good |
| Fish 1997 | Unclear | Unclear | Low | Low | Low | High | Unclear | Yes | Yes | Yes | Unclear | No | Poor |
| Green 2006 | Unclear | Unclear | Low | Low | Low | Unclear | Unclear | Yes | NA | Yes | Yes | No | Fair |
| Helenius 2003 | Unclear | Unclear | Low | Low | Low | Unclear | Unclear | Yes | NA | Unclear | Yes | No | Fair |
| Huang 2003 | Unclear | Unclear | Low | Low | Low | Low | Unclear | No | Yes | Yes | Unclear | No | Good |
| Israel 2002 | Low | Low | Low | Low | Low | Low | Low | No | Yes | Yes | Unclear | No | Good |
| Jayaram 2005 | Unclear | Unclear | Unclear | Unclear | Unclear | Low | Unclear | No | NA | Yes | Yes | No wash-out period | Poor |
| Jayaram 2005 | Unclear | Unclear | Low | Low | Low | Low | Unclear | No | Unclear | Yes | Yes | No | Good |
| Kanazawa 2004 | Unclear | Unclear | Unclear | Unclear | Unclear | Unclear | Unclear | Unclear | NA | Yes | Unclear | No | Fair |
| Kanniess 2002 | Unclear | Unclear | Low | Low | Low | Unclear | Unclear | Yes | Yes | Yes | Yes | Unclear | Fair |

Table 4. (continued)

| Source | Sequence generation | Allocation concealment | Blinding/ Patients | Blinding/ Caregivers | Blinding/ Assessors | Attrition | Selective outcome reporting | Intention-to-treat analysis | Baseline balance | Co-interventions similarity | Compliance | Other biases | Overall grade |
|------------------------|---------------------|------------------------|--------------------|----------------------|---------------------|-----------|-----------------------------|-----------------------------|------------------|-----------------------------|------------|---|---------------|
| Kraft 2006 | Unclear | Unclear | Unclear | Unclear | Unclear | Unclear | Unclear | Unsure | NA | Yes | Unsure | No | Fair |
| Laviolette 1999 | Low | Unclear | Low | Low | Low | Low | Low | No | Yes | Yes | Yes | No | Good |
| Leff 1998 | Unclear | Unclear | Low | Low | Low | Low | Unclear | No | Yes | Yes | Unsure | No | Good |
| Löfdahl 1999 | Low | Unclear | Unclear | Unclear | Unclear | High | Low | Unsure | Yes | No | Unsure | No | Poor |
| Malmstrom 1999 | Low | Unclear | Low | Low | Low | High | Low | No | Yes | Yes | Yes | No | Poor |
| Minoguchi 2002 | Unclear | Unclear | Low | Low | Low | Low | Unclear | No | NA | Yes | Unsure | No | Good |
| Nakamura 1998 | Unclear | Unclear | Low | Low | Low | Low | Unclear | No | No | Yes | Unsure | No | Poor |
| Nathan 1998 | Unclear | Unclear | Low | Low | Low | Low | Unclear | No | Yes | Yes | Unsure | No | Good |
| Nathan 1999 | Unclear | Unclear | Low | Low | Low | High | Unclear | No | Yes | Yes | Unsure | No | Poor |
| Nathan 2005 | Unclear | Unclear | Low | Low | Low | Low | Unclear | No | Yes | Yes | Unsure | No | Poor |
| Pizzichini 1999 | Low | Unclear | Low | Low | Low | Low | Unclear | Yes | Yes | Yes | Unsure | No | Good |
| Reid 2008 | Low | Low | Low | Low | Low | Low | Unclear | Unsure | No | Yes | Yes | No | Poor |
| Reiss 1998 | Low | Unclear | Low | Low | Low | High | Low | No | Yes | Yes | Unsure | No | Poor |
| Schäper 2011 | Unclear | Unclear | Unclear | High | High | Low | Unclear | Yes | NA | Unsure | Unsure | No wash-out period reported, differential duration of treatment periods | Poor |
| Spahn 2006 | Unclear | Unclear | Low | Low | Low | Low | Unclear | Yes | Unsure | Yes | Unsure | No | Fair |
| Spector 1994 | Unclear | Unclear | Unclear | Unclear | Unclear | Unclear | Unclear | No | Yes | Yes | Unsure | No | Fair |

Table 4. (continued)

| Source | Sequence generation | Allocation concealment | Blinding/ Patients | Blinding/ Caregivers | Blinding/ Assessors | Attrition | Selective outcome reporting | Intention-to-treat analysis | Baseline balance | Co-interventions similarity | Compliance | Other biases | Overall grade |
|-----------------------|---------------------|------------------------|--------------------|----------------------|---------------------|-----------|-----------------------------|-----------------------------|------------------|-----------------------------|------------|-----------------------------|---------------|
| Spector 2004 | Unclear | Unclear | Low | Low | Low | Low | Unclear | Unsure | No | Yes | Unsure | No | Fair |
| Stelmach 2007 | Low | Unclear | Low | Low | Low | Low | Unclear | No | Yes | Yes | Unsure | No | Good |
| Stelmach 2002 | Low | Unclear | Low | Low | Low | Low | Unclear | No | Yes | Yes | Unsure | No | Good |
| Stelmach 2002 | Low | Unclear | Unclear | Unclear | Unclear | Low | Unclear | No | Yes | Yes | Unsure | No | Fair |
| Storms 2004 | Low | Low | Low | Low | Low | Low | Unclear | No | Yes | Yes | Unsure | No | Good |
| Strunk 2008 | Unclear | Unclear | Low | Low | Low | High | Unclear | Yes | Unsure | No | Yes | Stopped early | Poor |
| Tohda 2002 | Unclear | Unclear | Low | Low | Low | Low | Unclear | No | Yes | No | Unsure | No | Good |
| Ulrik 2009 | Unclear | Unclear | Low | Low | Low | Low | Unclear | Yes | Yes | Yes | Yes | No | Good |
| Vaquerizo 2003 | Low | Unclear | Low | Low | Low | Low | Unclear | No | Yes | Yes | Unsure | No | Good |
| Wise 2009 | Low | Unclear | Low | Low | Low | Low | Unclear | No | Yes | Yes | Unsure | No | Good |
| Yoo 2001 | Low | Unclear | Low | Low | Low | Unclear | Unclear | No | Yes | Yes | Unsure | No | Fair |
| Yoshida 2002 | Unclear | Unclear | Low | Low | Low | Unclear | Unclear | Unsure | NA | Unsure | Unsure | No | Fair |
| Zeidler 2006 | Unclear | Unclear | Low | Low | Low | Unclear | Unclear | No | NA | Yes | Unsure | No wash-out period reported | Poor |

Abbreviations: NA: Non applicable

Table 5. List of included trials and contribution to meta-analysis of each outcome

| Source | Exacerbations | FEV ₁ | Daytime symptoms | SABA use | Nocturnal awakenings | Asthma specific quality of life |
|--------------------------------|--------------------------|--|------------------|----------------------------|----------------------|---------------------------------|
| Altman 1998 | | X (% change from baseline) | X | X (change from baseline) | | |
| American Lung Association 2007 | X | X (change from baseline in liters) | | | | X |
| Awad 2002 | | X (change from baseline in liters) X (change from baseline in FEV ₁ predicted) | X | | | |
| Baena-Cagnani 2003 | | | | X (% change from baseline) | | |
| Barnes 1997 | | X (change from baseline in liters) | X | X (change from baseline) | | |
| Baumgartner 2003 | X | X (% change from baseline) | | X (% change from baseline) | | |
| Busse 2001 | X | X (% change from baseline) | X | X (change from baseline) | X | |
| Cakmak 2004 | | | | | | |
| Dahlén 2002 | | X (% change from baseline) | X | X (change from baseline) | | X |
| Fish 1997 | | X (change from baseline in liters) X (change from baseline in FEV ₁ % predicted) | | X (change from baseline) | X | |
| Green 2006 | | | | | | |
| Helenius 2004 | | | | | | |
| Huang 2003 | | | | | | |
| Israel 2002 | X | X (change from baseline) | | X (% change from baseline) | | |
| Jayaram 2005 | | | | | | |
| Jayaram 2005 | | X (change from baseline) | X | X (change from baseline) | | |
| Kanazawa 2004 | | | | | | |
| Kanniess 2002 | | X (change from baseline) | X | X (change from baseline) | | |
| Kraft 2006 | | X (sensitivity analysis) | | | | |
| Laviolette 1999 | X | X (% change from baseline) | X | X (% change from baseline) | X | |
| Leff 1998 | | X (change from baseline) | | | | |
| Löfdahl 1999 | | X (% change from baseline) | | | | |
| Malmstrom 1999 | X | X (% change from baseline) | X | X (% change from baseline) | X | X |
| Minoguchi 2002 | | | | | | |
| Nakamura 1998 | | | | | | |
| Nathan 1998 | X (sensitivity analysis) | X (change from baseline) X (change from baseline in FEV ₁ % predicted) | | X (change from baseline) | X | |
| Nathan 1999 | | | | | | |
| Nathan 2005 | | | | | | |
| Pizzichini 1999 | | X (% change from baseline) | X | X (% change from baseline) | | |
| Reid (A) 2008 | | X (change from baseline) | | X (change from baseline) | | |
| Reid (B) | | | | | | |
| Reiss 1998 | X | X (% change from baseline) | X | X (% change from baseline) | X | |
| Schäper 2011 | | | | | | |

Table 5. (continued)

| Source | Exacerbations | FEV ₁ | Daytime symptoms | SABA use | Nocturnal awakenings | Asthma specific quality of life |
|-------------------|---------------|--|------------------|----------------------------|--------------------------|---------------------------------|
| Spahn 2006 | X | X (change from baseline in FEV ₁ % predicted) | X | X (change from baseline) | | |
| Spector 1994 | | X (% change from baseline) | | X (change from baseline) | X | |
| Spector 2004 | | | | | | |
| Stelmach (A) 2007 | | X (change from baseline in FEV ₁ % predicted) | | | | |
| Stelmach (B) | | | | | | |
| Stelmach 2002 | | X (change from baseline in FEV ₁ % predicted) | | | | |
| Stelmach 2002 | | X (change from baseline in FEV ₁ % predicted) | | | | |
| Storms 2004 | | X (change from baseline in FEV ₁ % predicted) | | | | |
| Strunk 2008 | | | | | | |
| Tohda 2002 | | | | | | |
| Ulrik 2009 | | | | | | |
| Vaquerizo 2003 | | X (% change from baseline) | X | X (% change from baseline) | | X |
| Wise (A) 2009 | X | X (change from baseline) | | | | X |
| Wise (B) | | | | | | |
| Yoo 2001 | | X (change from baseline) | X | | | |
| Yoshida 2002 | | | | | | |
| Zeidler 2006 | | X (sensitivity analysis) | | X (sensitivity analysis) | X (sensitivity analysis) | |

Studies not included in the meta-analyses

A summary of the results from the trials included in the systematic review, but not meta-analyzed, is presented in Table 6. The reasons for exclusion are also provided. Seven crossover trials were not included in the meta-analyses due to inadequate reporting or missing data.

Table 6. Effect of LTRAs versus placebo in trials not included in the meta-analyses

| Outcome measure | Study | LTRA, (95% CI) | Placebo (95% CI) | Reported p-value (difference in effects) | Reason for exclusion |
|---|-----------------------------|---------------------|---------------------|---|---|
| FEV₁(L) change from baseline | | | | | |
| | Baena-Cagnani 2003 | 0.18 | 0.05 | p < 0.01 | data provided only for patients with FEV ₁ % predicted less than 80% |
| | Cakmak 2004 | ND | ND | NS | no effect estimates reported |
| | Green 2006 [□] | -0.05 (-0.13, 0.03) | -0.07 (-0.20, 0.05) | NS | crossover trial |
| | Helenius 2004 | -0.01 | 0.02 | NS | crossover trial |
| | Jayaram 2005 [□] | 0.09 | 0.02 | NS | crossover trial |
| | Nathan 1999 | 0.23 | 0.04 | p < 0.05 | patients with PEF ≥ 10% |
| | | 0.15 | 0.07 | NS | patients with PEF < 10% |
| | Schäper 2011 [‡] | 0.2 | -0.5 | NS overall | 1 st period of crossover |
| | | 0.2 | -0.2 | | 2 nd period of crossover |
| | Spector 2004 | ND | ND | NS | no effect estimates reported |
| | Tohda 2002 | ND | ND | NS | no effect estimates reported |
| | Ulrik 2009 | ND | ND | NS | no effect estimates reported |
| | Yoshida 2002 [□] | 9.8 | -0.2 | p < 0.05 | crossover trial |
| FEV₁ % predicted change from baseline | | | | | |
| | Helenius 2004 | -1.5 | -1.3 | NS | crossover trial |
| | Jayaram 2005 [□] | 3.6 | 1.5 | p = 0.4 | crossover trial |
| | Nakamura 1998 | -3.5 median | 1.1 median | NS | data reported as medians |
| | Schäper 2011 [‡] | 3.6 | -0.3 | NS overall | 1 st period of crossover |
| | | 5 | -1.7 | | 2 nd period of crossover |
| Daytime symptoms change from baseline | | | | | |
| scale: 0-3 (more symptoms) | Fish 1997 [†] | - | - | p < 0.01 (-1.4) | expressed as weekly totals |
| scale: 0-3 (more symptoms) | Green 2006 [□] | -0.09 (-0.27, 0.10) | -0.07 (-0.20, 0.06) | NS | crossover trial |
| scale: 5-35 (fewer symptoms) | Jayaram 2005 [□] | -0.3 | 0.6 | p = 0.6 | crossover trial |
| scale: 0-6 (more symptoms) | Löfdahl 1999 | 0.07 | 0.12 | NS | inadequate reporting |
| scale: 0.5-10 (more symptoms) | Minoguchi 2002 [□] | -1.6 | 0 | - | crossover trial |
| scale: 0-3 (more symptoms) | Nathan 1998 [†] | - | - | p < 0.01 (-0.14) | 'adjusted' treatment effect scale: 0-3 (more symptoms) |
| scale: 0-3 (more symptoms) | Spector 1994 | -27% | -13% | p ≤ 0.01 | reported as percentages |
| scale: 0-4 (more symptoms) | Zeidler 2006 [*] | -0.2 | 0.2 | NS | crossover trial |

Table 6. (continued)

| Outcome measure | Study | LTRA, (95% CI) | Placebo, (95% CI) | Reported p-value (difference in effects) | Reason for exclusion |
|---|---------------------------|--------------------|----------------------|---|--|
| SABA use (puffs/day) change from baseline | | | | | |
| | Awad 2002 | -14.3 | -0.6 | NS | number of SABA tablets |
| | Cakmak 2004 | ND | ND | NS | no effect estimates reported |
| | Green 2006 [□] | -0.6 (-1.14, 0.06) | 0.9 (-0.07, 2.5) | NS | crossover trial |
| | Jayaram 2005 [□] | -16 | -9.4 | NS | µg/day |
| | Nakamura 1998 | 1.4 median | 0.3 median | p < 0.01 (-1.4 median) | data reported as medians |
| | Nathan 1999 | -1.79 | 0.1 | p < 0.05 | patients with PEF ≥ 10% |
| | | -0.75 | -0.31 | NS | patients with PEF < 10% |
| | Schäper 2011 [‡] | ND | ND | p < 0.05 | crossover trial |
| | Yoo 2001 | -5.36 | 0.28 | - | puffs/2 weeks |
| Nocturnal awakenings (per week) change from baseline | | | | | |
| | Altman 1998 | -1.4 (-2.1,-0.8) | 0.8 (-1.4,-0.1) | - | N analyzed missing |
| | Dahlén 2002 | -0.18 | -0.04 | - | N analyzed missing |
| | Nathan 1999 | -1 | -0.23 | p < 0.05 | patients with PEF ≥ 10% |
| | | -0.7 | -0.37 | NS | patients with PEF < 10% |
| | Nathan 2005 | ND | ND | NS | no effect estimates reported |
| | Reid (A) 2008 | 0.2 | 0.7 | - | change in total awakenings |
| | Reid (B) 2008 | -1.6 | -1.1 | - | change in total awakenings |
| | Wise (A) 2009 | - | - | NS | patients with at least 1 awakening |
| | Wise (B) 2009 | - | - | NS | patients with at least 1 awakening |
| | Zeidler 2006* | -0.1 | 0 | NS | |
| ASQL change from baseline | | | | | |
| scale: 0-6 (better) | Altman 2002 | - | - | - | pooled effect for all doses used in the trial (scale 0-6) |
| scale: 1-7 (better) | Busse 2001 | ND | ND | - | no clinically meaningful difference for any AQLQ domain |
| scale: 1-7 (better) | Green 2006 [□] | 0.2 (0,0.2) | 0 (-0.2,0.3) | NS | crossover trial |
| scale: ND | Spector 2004 | ND | ND | - | significant improvement |
| scale: 1-7 (better) | Zeidler 2006* | 0.5 | 0.1 | p = 0.04 | |

Abbreviations: ASQL: Asthma-specific quality of life; AQLQ: Asthma quality of life questionnaire; CI: Confidence interval; FEV₁: Forced expiratory volume in one second; ND: No data; NS: Not statistically significant; PEF: Peak Expiratory Flow; SABA: Short acting β₂-agonist use.

* Included in sensitivity analyses for FEV₁, SABA use, and nocturnal awakenings.

† Not meta-analyzed because final values and changes from baseline should not be combined together as standardized mean differences⁷⁴

‡ Not meta-analyzed because timing of assessment of endpoints differed between comparator groups (6 weeks vs. 4 weeks)

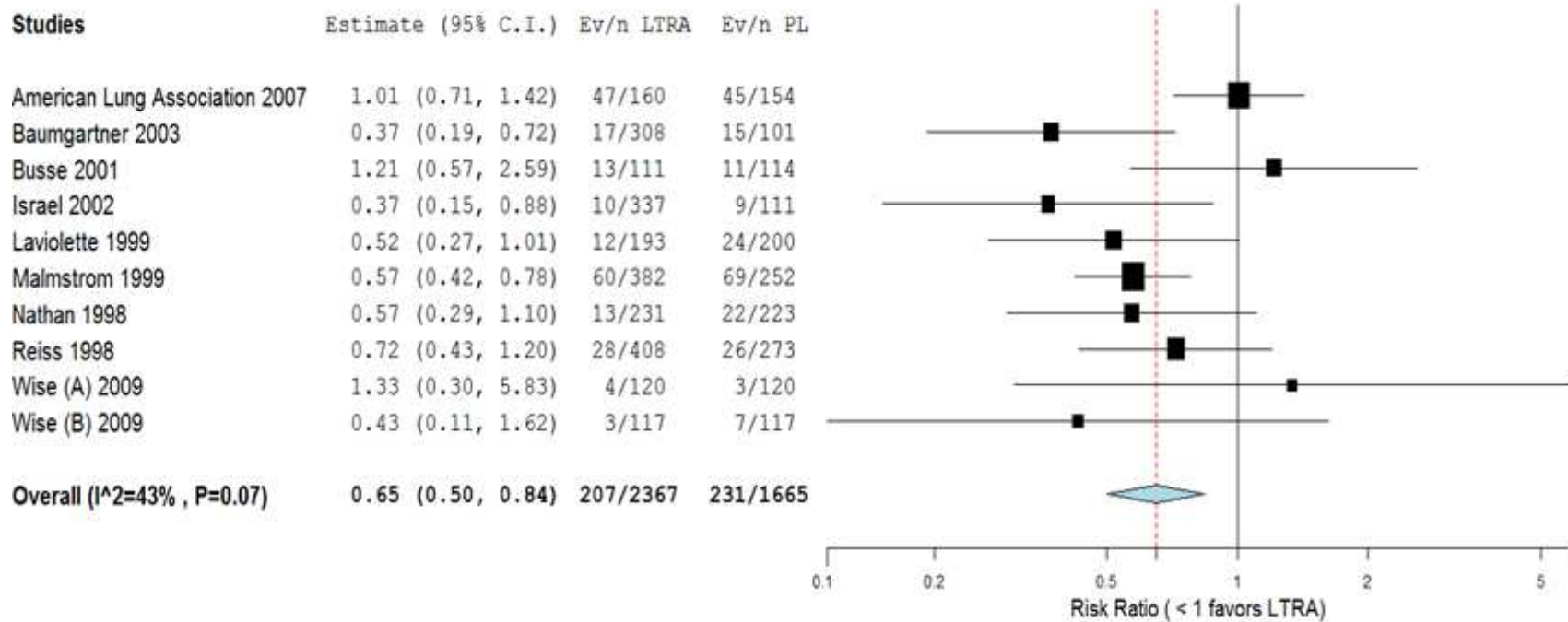
□ Not meta-analyzed due to missing correlation

Asthma exacerbations

The definition of an exacerbation reported in nine RCTs was consistent with our definition and, therefore, these trials were included in our main analysis (one RCT included two separate comparisons of LTRA vs. placebo in different patients). All definitions of exacerbations reported in the included RCTs are presented in Table 7. The main analysis included 2367 patients who received a LTRA and 1665 patients who received placebo.

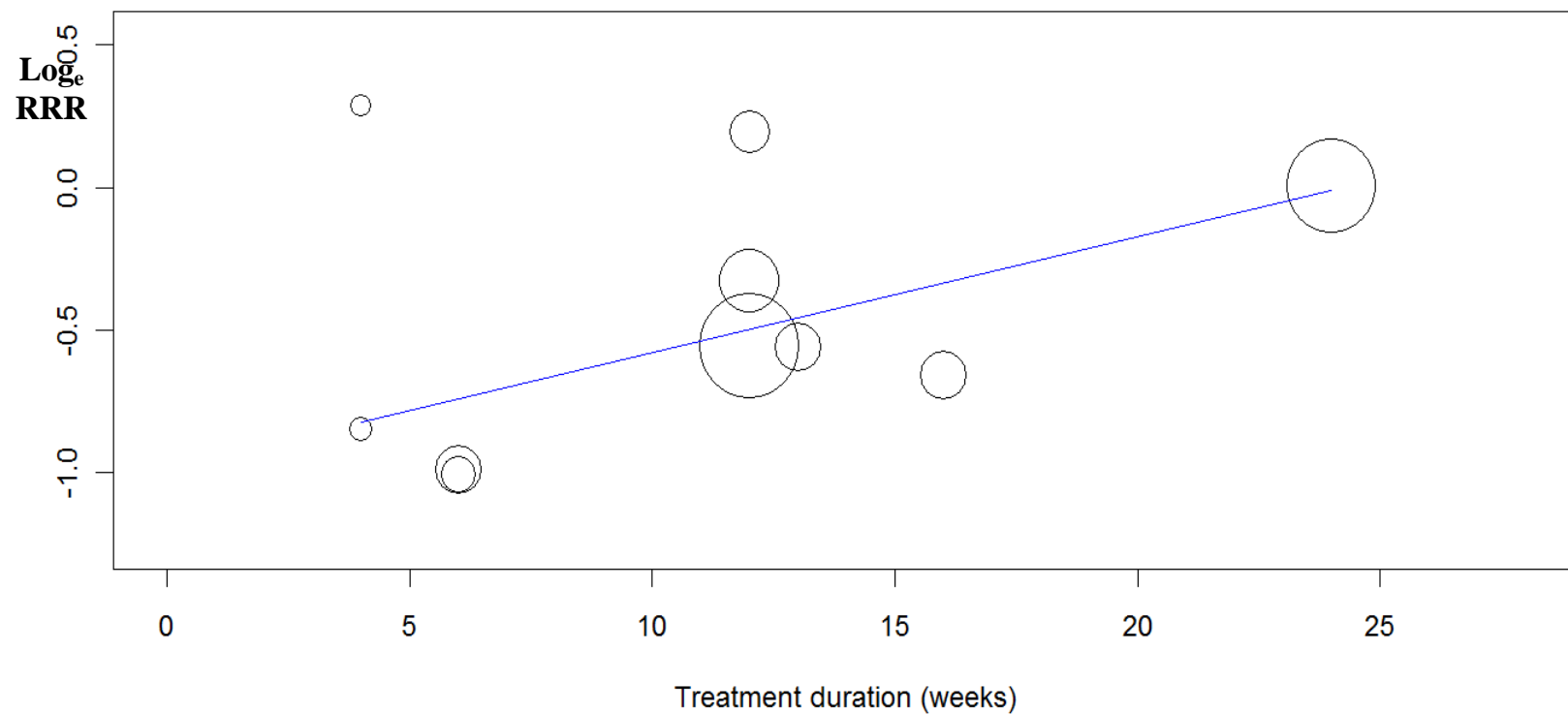
Overall, patients treated with LTRAs displayed a 35% decreased risk of experiencing an exacerbation compared to those treated with placebo (summary RR = 0.65, 95% CI: 0.5, 0.84) (Figure 2). The observed statistical heterogeneity was moderate ($I^2 = 43\%$). The addition of the one RCT with an unclear definition of an exacerbation did not change the results (summary RR = 0.65, 95% CI: 0.51, 0.84, $I^2 = 37\%$). The results of the univariate meta-regressions that we used in order to explore the effect of study-level characteristics on the pooled treatment effect are presented in Table 8. Across studies, the magnitude of the effect appeared to weaken as the study duration increased; the summary RR increased by 4% with every additional week of treatment duration (Relative Risk Ratio = 1.04, 95% CI: 1.01, 1.07) [Figure 3]. The limited number of studies in which all patients used ICS precluded meta-regression of this factor. Similarly, we could not assess the impact of the type of LTRA on the summary estimate because montelukast was administered in seven RCTs and zafirlukast in two. We could not examine the effect of allergic rhinitis or aspirin-induced asthma on the summary estimate due to incomplete reporting in the RCTs included.

Figure 2. Summary forest plot for asthma exacerbation



The center of the diamond (red dotted line) represents the pooled Risk Ratio and its size the length of the 95% Confidence Interval. Risk Ratios (squares) and 95% CIs (horizontal lines) for individual studies are also shown. The size of the squares is proportional to the weight of each study in the meta-analysis.

Figure 3. Meta-regression plot of \log_e relative risk ratio for exacerbation by treatment duration



Blue line represents the change in \log_e relative risk ratio. The relative risk ratio is defined as the ratio of Risk Ratios for exacerbation for each 1 week change in treatment duration. Circles represent studies included in the analysis and their size is proportional to the weight assigned in the meta-regression.

Table 7. Definitions of exacerbations in included trials

| Source | Reported definitions |
|---------------------------------------|---|
| Altman 1998 | more than 20% decrease in PEFR compared to baseline, more than 70% increase in SABA use, more than 50% increase in symptom score, “awake all night”, or unscheduled visit to a doctor or hospital (reported as percent of days with exacerbations) |
| American Lung Association 1997 | required use of oral corticosteroids or unscheduled health care (included in main analysis) |
| Awad 2002 | treated with ICS |
| Baumgartner 2003 | required use of oral corticosteroids or unscheduled medical care (included in main analysis) |
| Busse 2001 | required use of oral corticosteroids (included in main analysis) |
| Dahlen 2002 | more than 20% decrease in PEFR compared to baseline, more than 70% increase in SABA use, more than 50% increase in symptom score, “awake all night”, or unscheduled visit to a doctor or hospital (reported as percent of days with exacerbations) [4 exacerbations occurred in each group] |
| Green 2006 | required use of oral corticosteroids (Not included in main analysis because only number of events were reported) |
| Israel 2002 | required unscheduled visit to the doctor’s office or emergency department, hospitalization, or treatment with oral corticosteroids (included in main analysis) |
| Jayaram 2005 | treated with ICS |
| Kanniess 2002 | more than 50% decrease in PEFR compared to values at entry, or an increase in daytime symptoms of 3 or more, or in night-time symptoms of 2 or more on at least 3 consecutive days |
| Laviolette 1999 | required unscheduled visit, hospitalization, or treatment with oral corticosteroids (included in main analysis) |
| Leff 1998 | treated with ICS |
| Malmstrom 1999 | required unscheduled visit, hospitalization, or treatment with oral corticosteroids (included in main analysis) |
| Nathan 2005 | required treatment with asthma medications beyond study medications (included in sensitivity analysis) |
| Reiss 1998 | required oral corticosteroids (included in main analysis) |
| Spahn 2006 | required oral corticosteroids (included in main analysis) |
| Spector 1994 | no specific treatment protocol |
| Vaquerizo 2003 | more than 20% decrease in PEFR compared to baseline, more than 70% increase in SABA use, more than 50% increase in symptom score, “awake all night”, or unscheduled visit to a doctor or hospital (reported as percent of days with exacerbations) |
| Wise 2009 | required urgent asthma care or oral corticosteroids (included in main analysis) |
| Zeidler 2006 | required oral corticosteroids (not included in analysis because it is was not specified in which group of patients 2 exacerbations occurred) |

Abbreviations: ICS: Inhaled corticosteroids; PEFR: Peak expiratory flow rate; SABA: Short-acting β_2 -agonist.

Table 8. Meta-regression analyses examining the association of pre-specified covariates with the pooled treatment effect for exacerbation

| Covariate | Relative risk | 95% CI |
|------------------------|----------------|-----------|
| Treatment duration | 1.04 | 1.01-1.07 |
| Equal ICS use* | Not performed† | |
| ICS dose | Not performed† | |
| Type of LTRA | Not performed† | |
| Allergic rhinitis | Not performed‡ | |
| Aspirin-induced asthma | Not performed‡ | |

* No use of ICS was used as the reference group

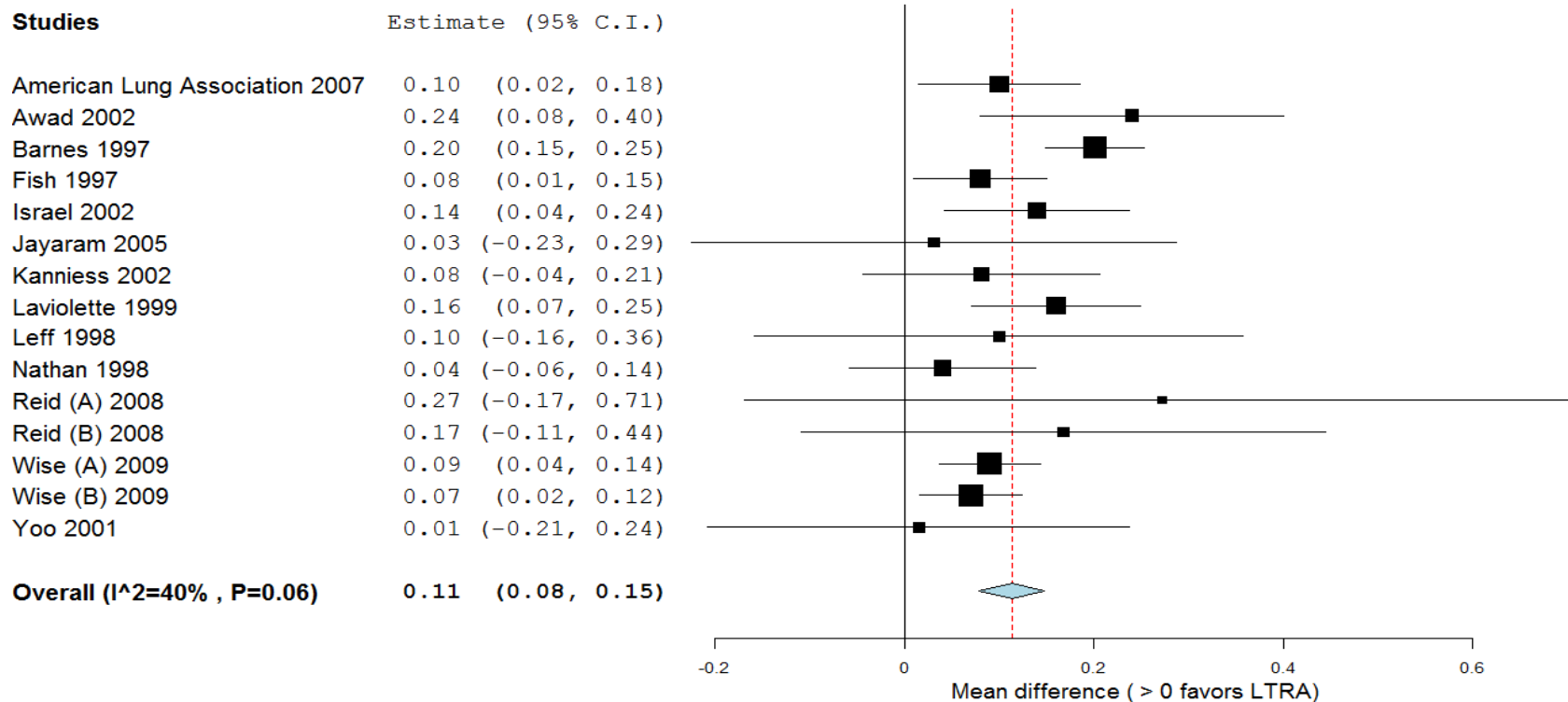
† Analyses were not performed due to insufficient number of RCTs in each subgroup

‡ Analyses were not performed due to inadequate reporting in the RCTs

Forced Expiratory Volume in one second

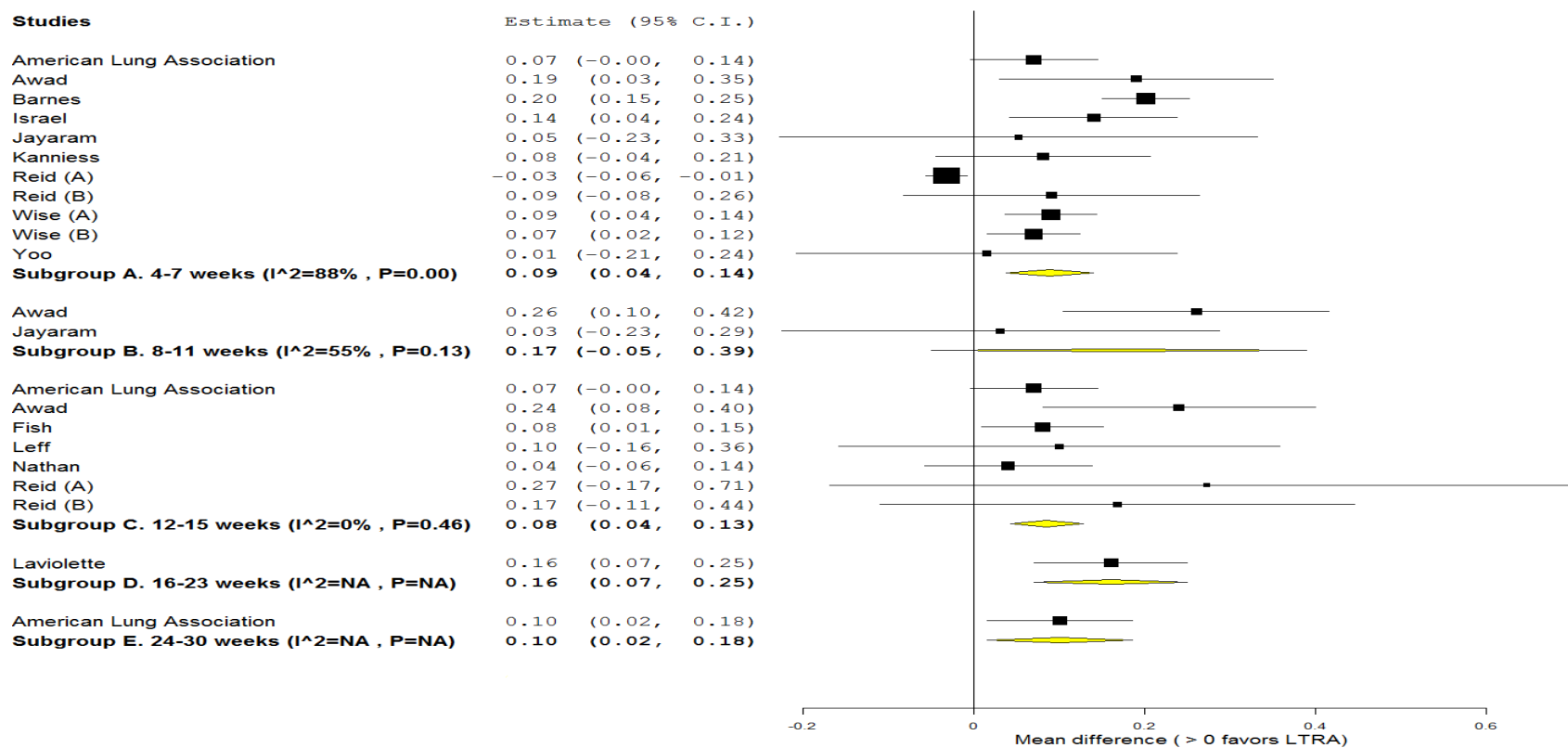
Forty-three trials reported the effect of LTRAs compared with placebo on FEV₁. Thirty-five trials reported FEV₁ in liters (L) and 15 as percent of predicted values (FEV₁ % predicted). A mean difference in change from baseline in FEV₁ (L) was computed for 13 trials, a mean difference in percent change from baseline was computed for 11 trials, and a mean difference in FEV₁ % predicted was computed for 8 trials. Overall, LTRAs significantly improved FEV₁ (L) compared to placebo (summary MD = 0.11, 95% CI: 0.08, 0.15) (Fig 4). The observed statistical heterogeneity was moderate ($I^2 = 48\%$). Using meta-regression, there was no association between treatment duration and the effect (p-value = 0.93) [Fig 5 and 6]. No association was observed between concomitant ICS use and the pooled effect (p-value = 0.39) [Fig 7]. There was no significant difference in the effect of zafirlukast compared to montelukast on the summary effect (p-value = 0.88) [Fig 8]. Results did not change with the inclusion of one crossover study (summary MD = 0.12, 95% CI: 0.08, 0.15).

Figure 4. Summary forest plot of mean difference for FEV₁ (L)



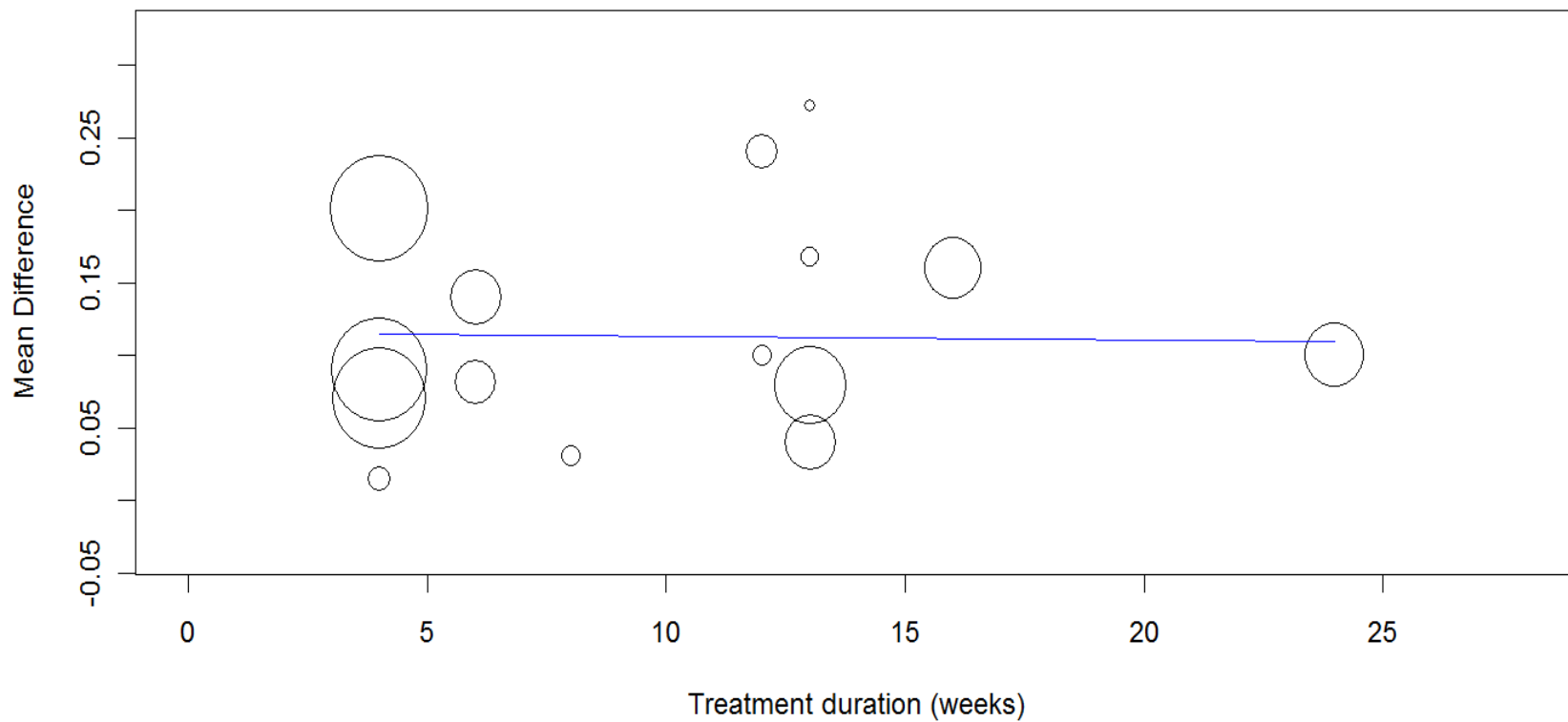
The center of the diamond (red dotted line) represents the pooled mean difference and its size the length of the 95% Confidence Interval. Mean differences (squares) and 95% CIs (horizontal lines) for individual studies are also shown. The size of the squares is proportional to the weight of each study in the meta-analysis.

Figure 5. Summary forest plot of mean difference by subgroups of treatment duration for FEV₁ (L)



The center of the yellow diamonds represents the pooled mean difference and its size the length of the 95% Confidence Interval in each subgroup. Mean differences (squares) and 95% CIs (horizontal lines) for individual studies are also shown. The size of the squares is proportional to the weight of each study in the meta-analysis. Some studies may contribute data to more than one subgroup.

Figure 6. Meta-regression plot of mean difference in FEV₁ by treatment duration



Blue line represents the change in mean difference. Circles represent studies included in the analysis and their size is proportional to the weights assigned in meta-regression.

Figure 7. Summary forest plot of mean difference by subgroups of ICS use for FEV₁ (L)

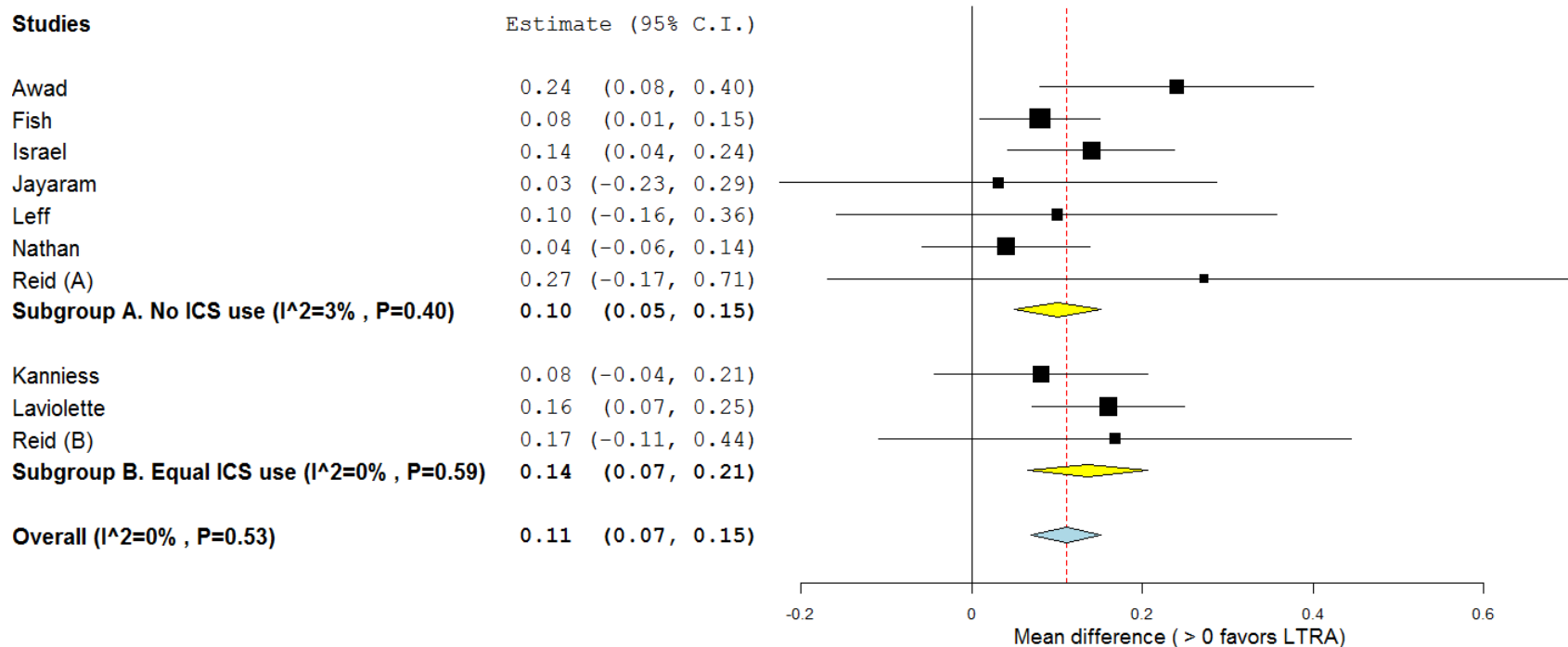
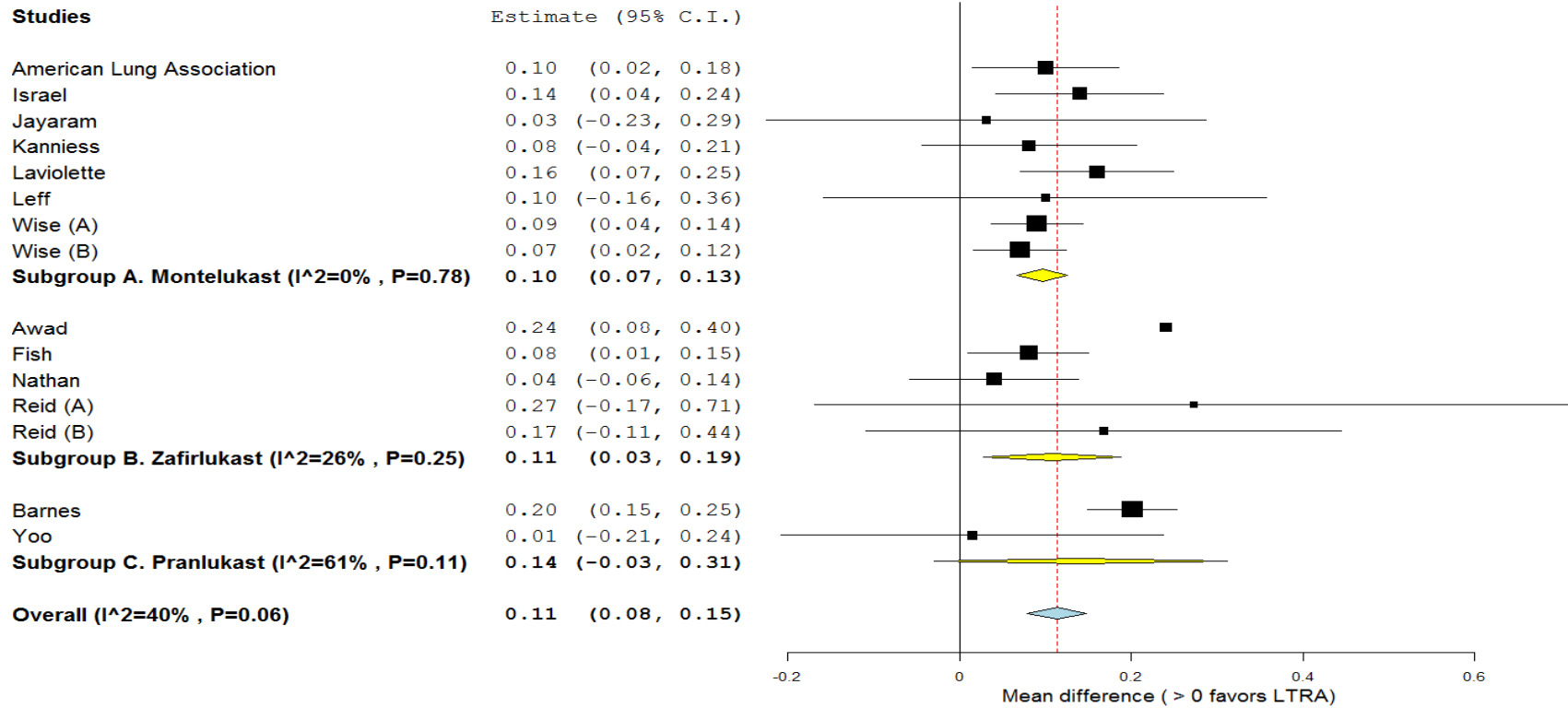


Figure 8. Summary forest plot of mean difference by subgroups of LTRA for FEV₁ (L)



The mean difference in percent change from baseline between LTRAs and placebo was statistically significant (summary MD = 5.95, 95% CI: 3.30, 8.60). Trials were heterogeneous ($I^2 = 69\%$) [Fig 9]. The treatment effect was larger in studies of shorter duration (MD = -0.62, 95% CI: -1.16, -0.08) [Fig 10 and 11]. The limited number of studies precluded meta-regression of other pre-specified factors (Fig 12 and 13).

LTRAs significantly improved FEV₁ % predicted (summary MD = 4.16, 95 % CI: 1.47, 6.85) [Fig 14]. Studies were significantly heterogeneous ($I^2 = 63\%$). No significant association was observed between treatment duration, concomitant ICS use, or type of LTRA with the pooled effect (p-value = 0.8, 0.4, 0.6, respectively) [Fig 15-18]. The summary estimate increased when two crossover trials were included in the meta-analysis (summary MD = 5.07, 95% CI: 2.46, 7.69).

Figure 9. Summary forest plot of mean difference in percent change from baseline

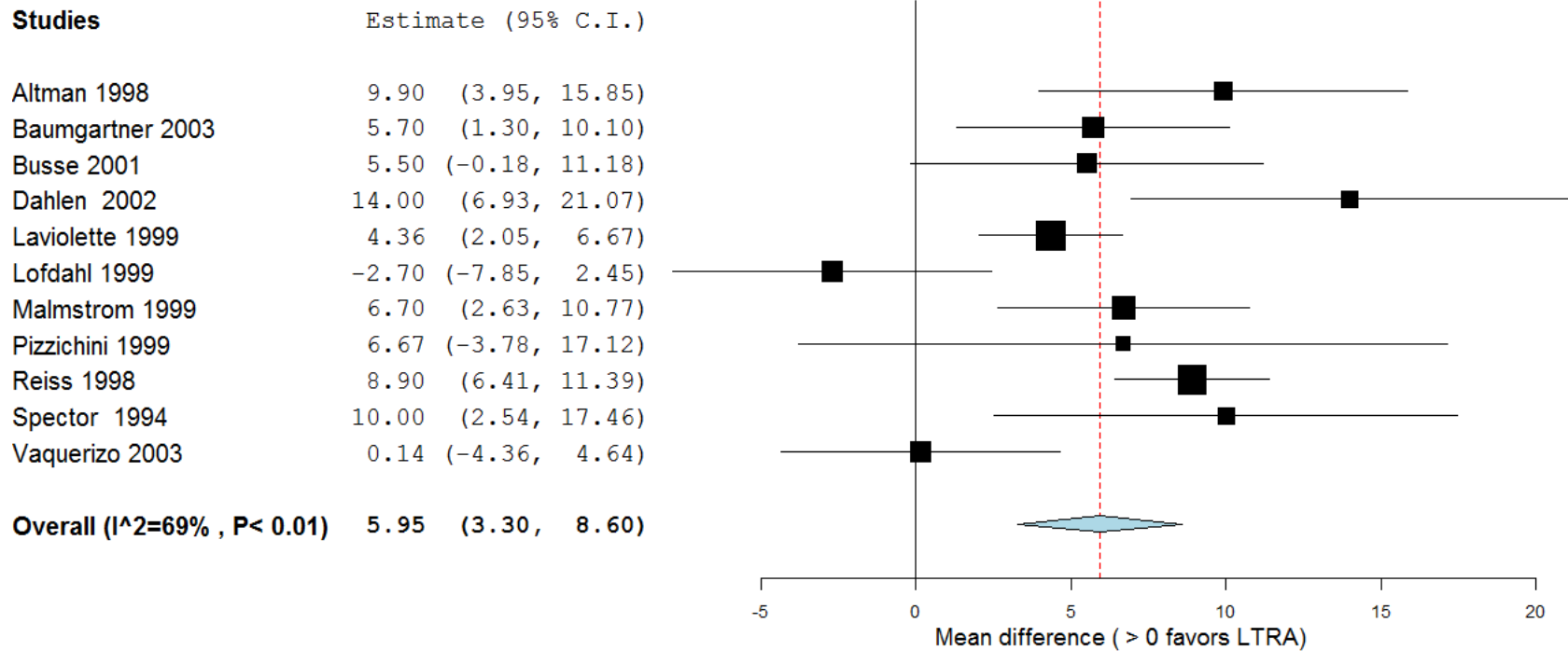


Figure 10. Meta-regression plot of mean difference in percent change from baseline for FEV₁ by treatment duration

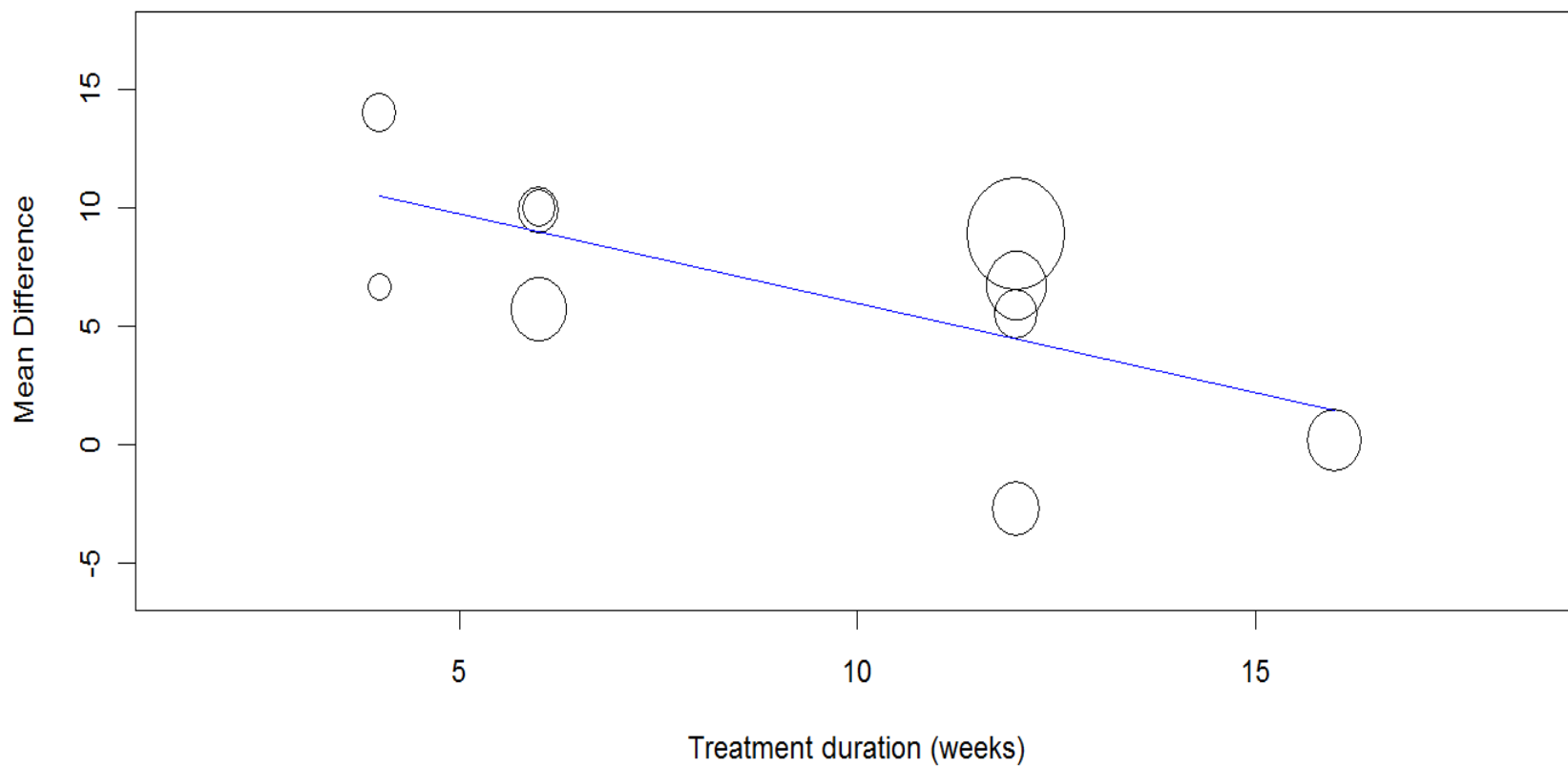


Figure 11. Summary forest plot of mean difference in percent change from baseline by subgroups of treatment duration for FEV₁

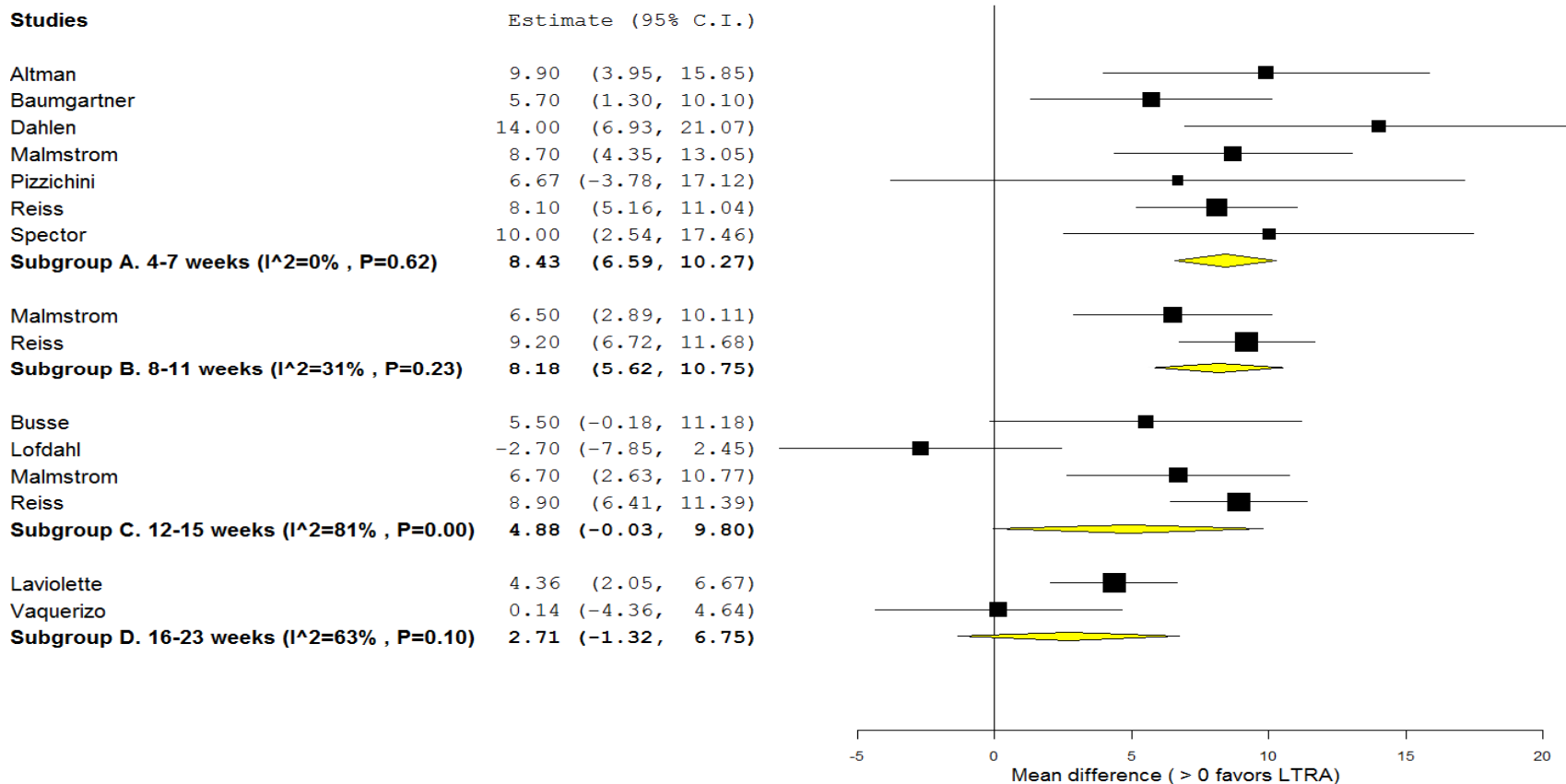


Figure 12. Summary forest plot of mean difference in percent change from baseline by subgroups of ICS use for FEV₁

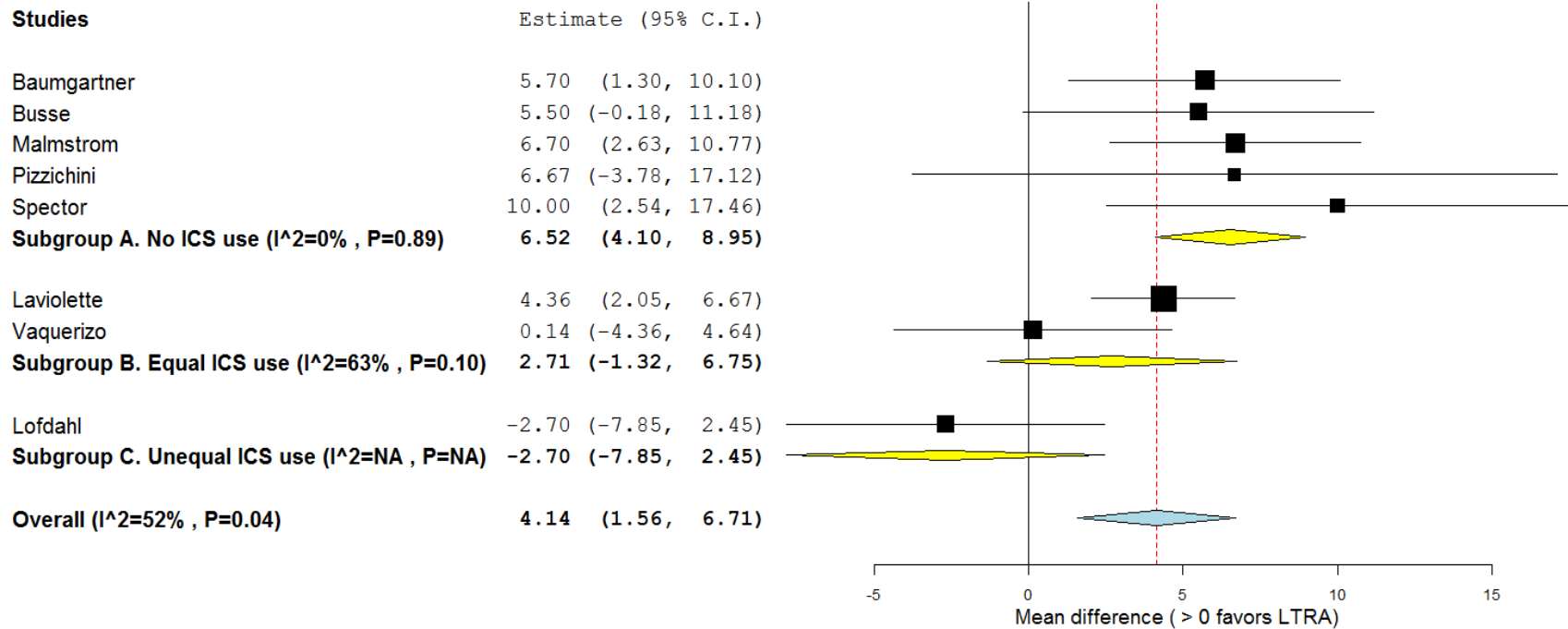


Figure 13. Summary forest plot of mean difference in percent change from baseline by subgroups of LTRAs for FEV₁

| Studies | Estimate (95% C.I.) |
|--|---------------------------|
| Baumgartner | 5.70 (1.30, 10.10) |
| Malmstrom | 6.70 (2.63, 10.77) |
| Pizzichini | 6.67 (-3.78, 17.12) |
| Altman | 9.90 (3.95, 15.85) |
| Dahlen | 14.00 (6.93, 21.07) |
| Laviolette | 4.36 (2.05, 6.67) |
| Reiss | 8.90 (6.41, 11.39) |
| Vaquerizo | 0.14 (-4.36, 4.64) |
| Lofdahl | -2.70 (-7.85, 2.45) |
| Subgroup Montelukast (I²=74% , P=0.00) | 5.70 (2.58, 8.82) |
| Busse | 5.50 (-0.18, 11.18) |
| Spector | 10.00 (2.54, 17.46) |
| Subgroup Zafirlukast (I²=0% , P=0.35) | 7.15 (2.63, 11.67) |
| Overall (I²=69% , P=0.00) | 5.95 (3.30, 8.60) |

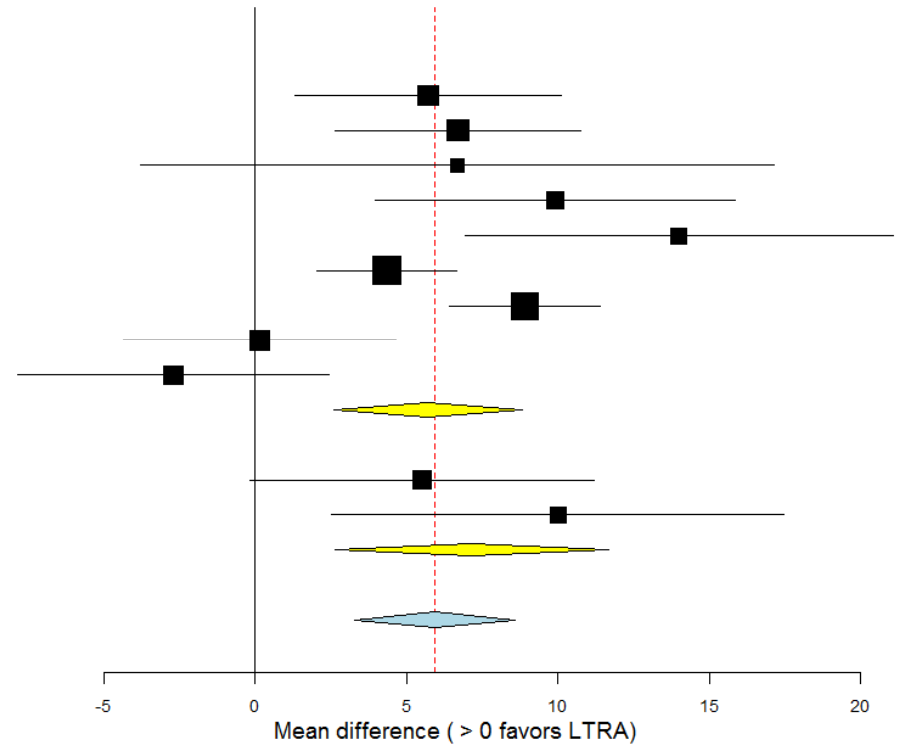


Figure 14. Summary forest plot of mean difference for FEV₁ % predicted

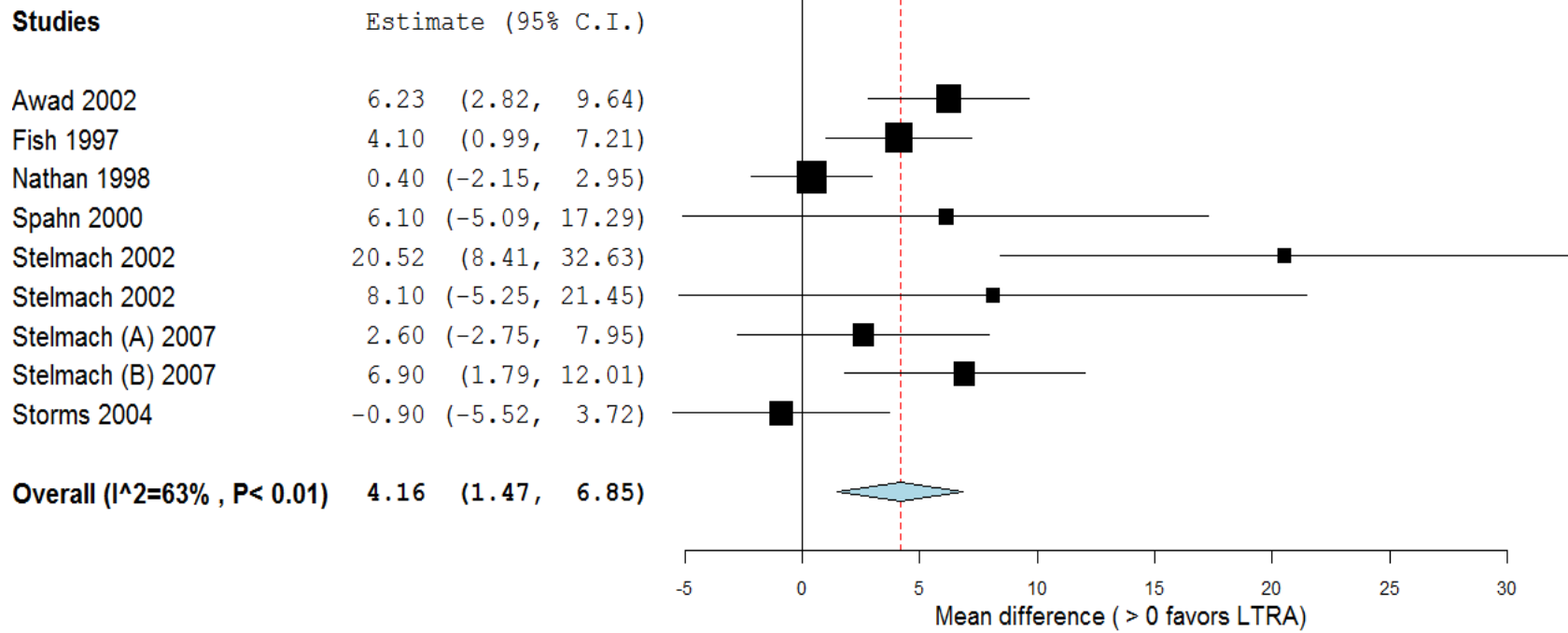


Figure 15. Meta-regression plot of mean difference for FEV₁ % predicted by treatment duration

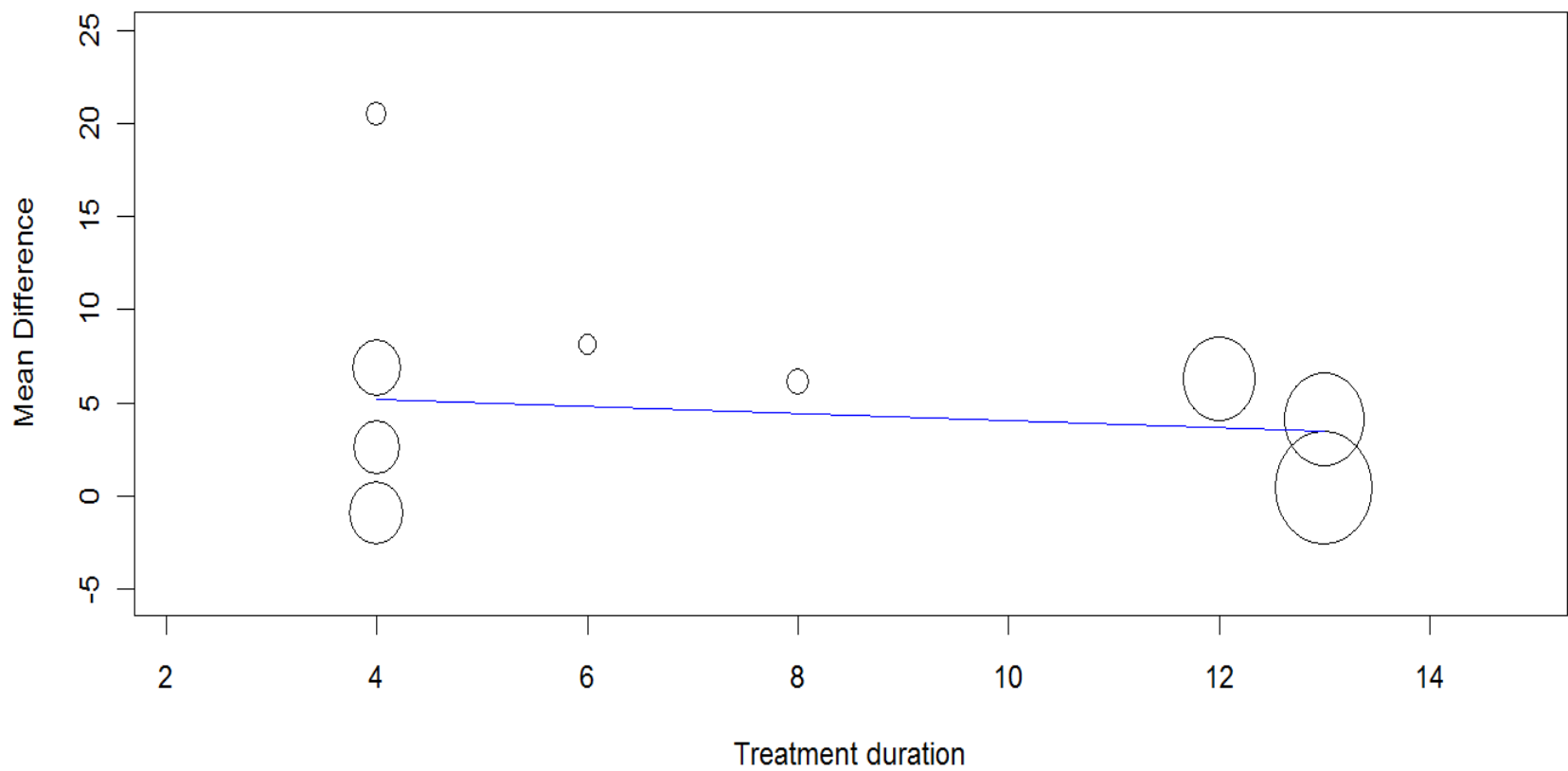


Figure 16. Summary forest plot of mean difference by subgroups of treatment duration for FEV₁ % predicted

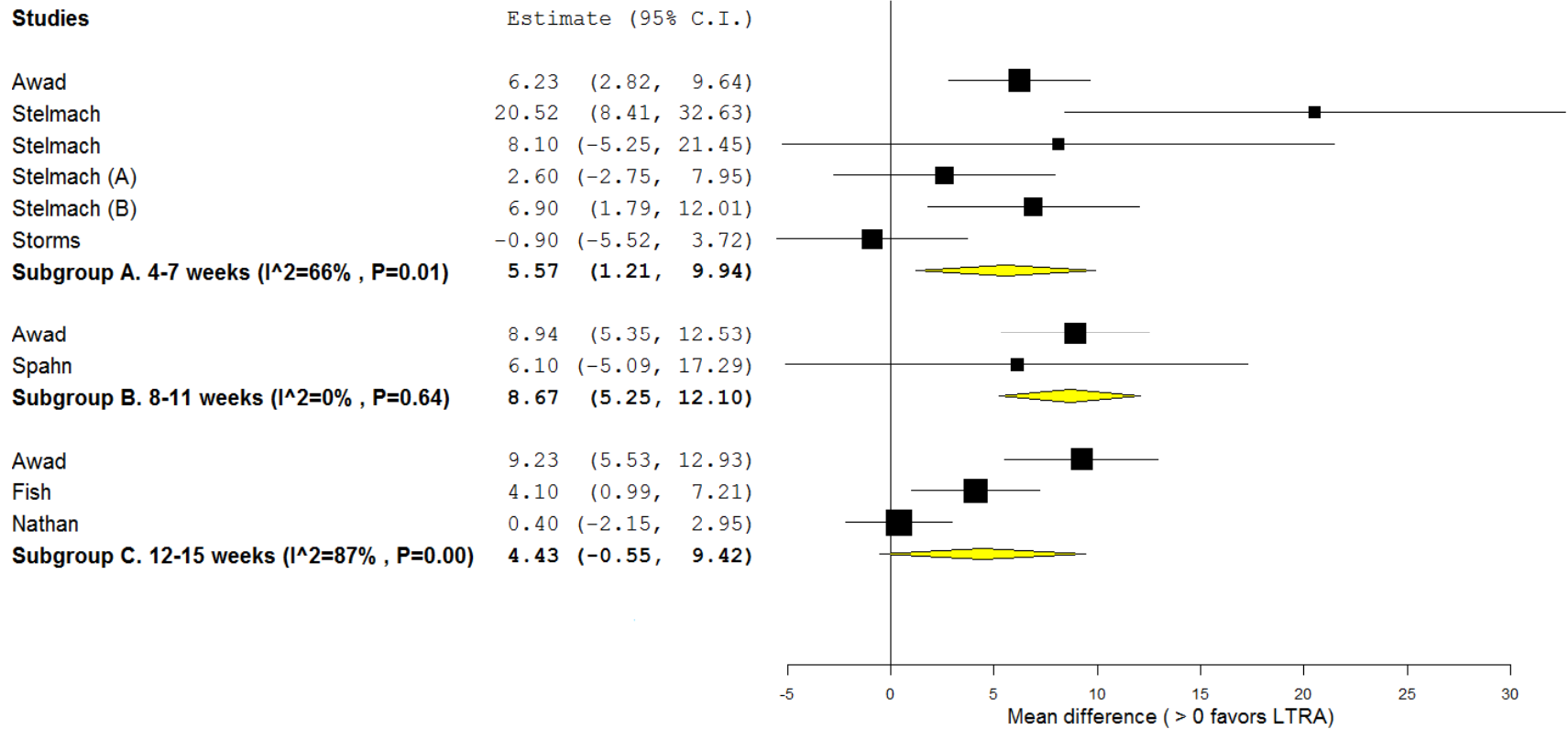


Figure 17. Summary forest plot of mean difference by subgroups of ICS use for FEV₁ % predicted

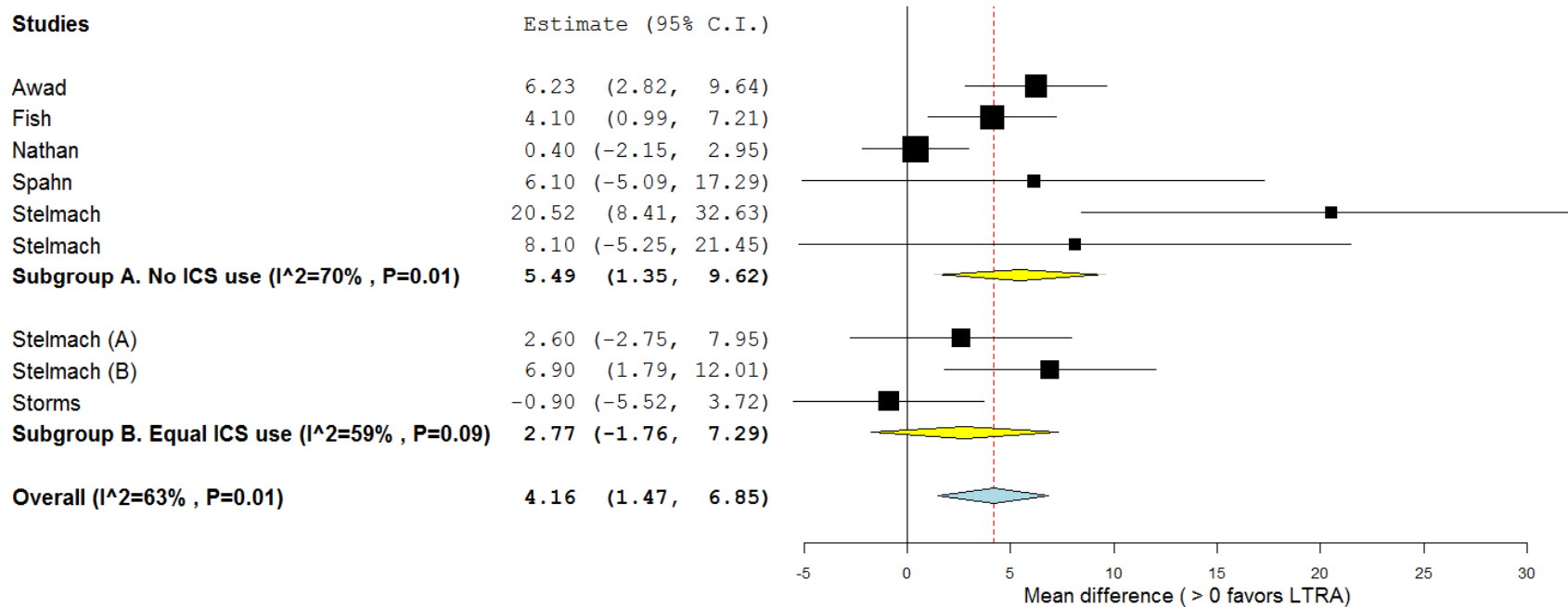
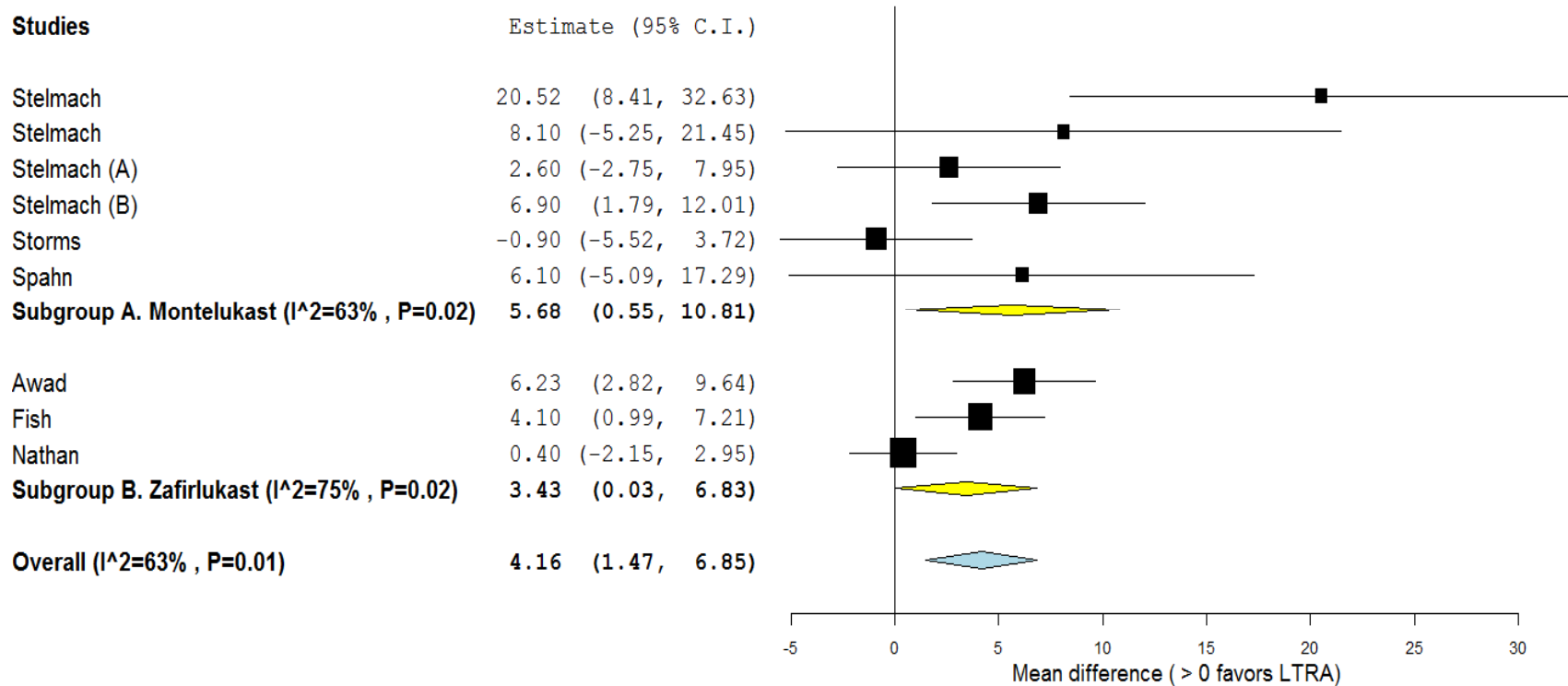


Figure 18. Summary forest plot of mean difference by subgroups of LTRA for FEV₁ % predicted



Daytime symptoms

Daytime symptom scores were reported in 22 RCTs. Fourteen trials contributed data to the meta-analysis. A variety of scales were used with lower values indicating fewer symptoms. LTRAs significantly reduced daytime symptoms (summary SMD = -0.21, 95%: -0.37, -0.04) [Fig 19]. The level of statistical heterogeneity was high ($I^2 = 70\%$). Treatment duration and concomitant ICS use were not associated with the pooled effect (p-value = 0.8, p-value = 0.43, respectively) [Fig 20-22].

Short acting β_2 -agonist use

The effect of LTRAs on SABA use compared to placebo was examined in 29 RCTs. We could compute a mean difference in change from baseline in SABA use for 11 trials and a mean difference in percent change from baseline for 8 trials. LTRAs decreased the number of inhalations per day by 0.35 (summary MD = -0.65, 95% CI: -0.82, -0.49). The level of statistical heterogeneity observed was low ($I^2 = 7\%$). [Fig 23]. The effect was consistent across time-points (p-value = 0.99) and types of LTRA [Fig 24-26]. Results did not change when one crossover trial was included (summary MD = -0.66, 95% CI: -0.82,-0.49).

The mean difference in percent change from baseline was statistically significant (summary MD = -16.39, 95% CI: -22.37, -10.41). High statistical heterogeneity was observed ($I^2 = 74\%$). No association with treatment duration was observed (p-value = 0.78) [Fig 27-28].

Nocturnal awakenings

Nocturnal awakenings were reported in 15 RCTs. Only seven RCTs could be meta-analyzed due to missing data in the original trials (primarily, missing number of participants analyzed). Overall, LTRAs reduced nocturnal awakenings per week (summary MD = -0.66, 95% CI: -1.01, -0.32). The statistical heterogeneity detected was high ($I^2 = 85\%$). [Fig 29]. Inclusion of one crossover trial did not affect the results (summary MD = -0.65, 95% CI: -0.89,-0.30). We did not use meta-regression due to the limited number of studies.

Asthma-specific quality of life

Eleven trials examined the effect of LTRAs compared to placebo on asthma-specific quality of life. However, only five trials contributed data to our meta-analysis. The main reason for missing data was reporting of either non-significant results or results were presented per quality domain. LTRAs improved quality of life (summary SMD = 0.13, 95% CI: 0.02, 0.23). Statistical heterogeneity was low ($I^2 = 23\%$) [Fig 30]. We did not use meta-regression due to the small number of studies.

Figure 19. Summary forest plot of standardized mean difference for daytime symptoms

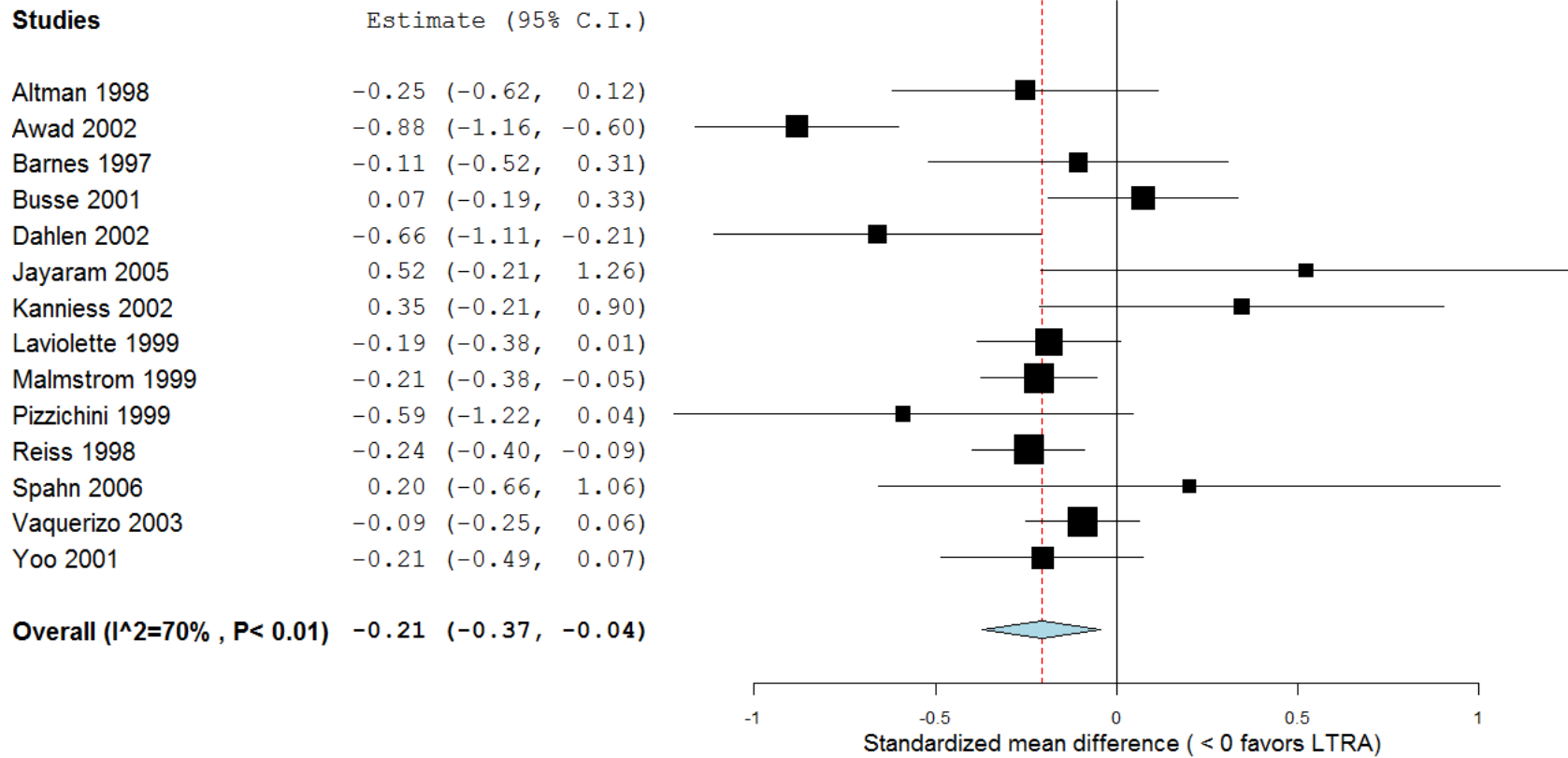


Figure 20. Meta-regression plot of standardized mean difference for daytime symptoms by treatment duration

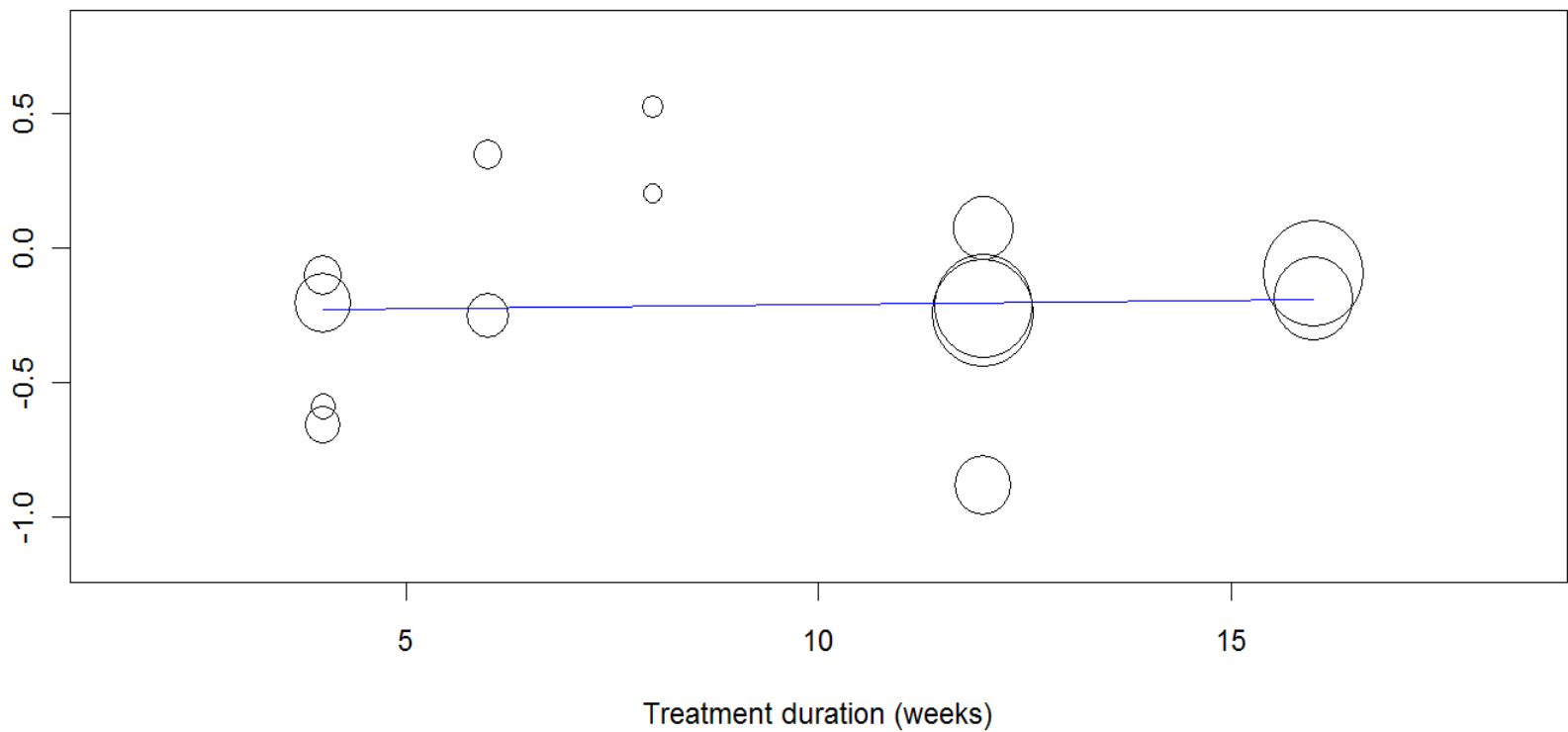


Figure 21. Summary forest plot of standardized mean difference by subgroups of treatment duration for daytime symptoms

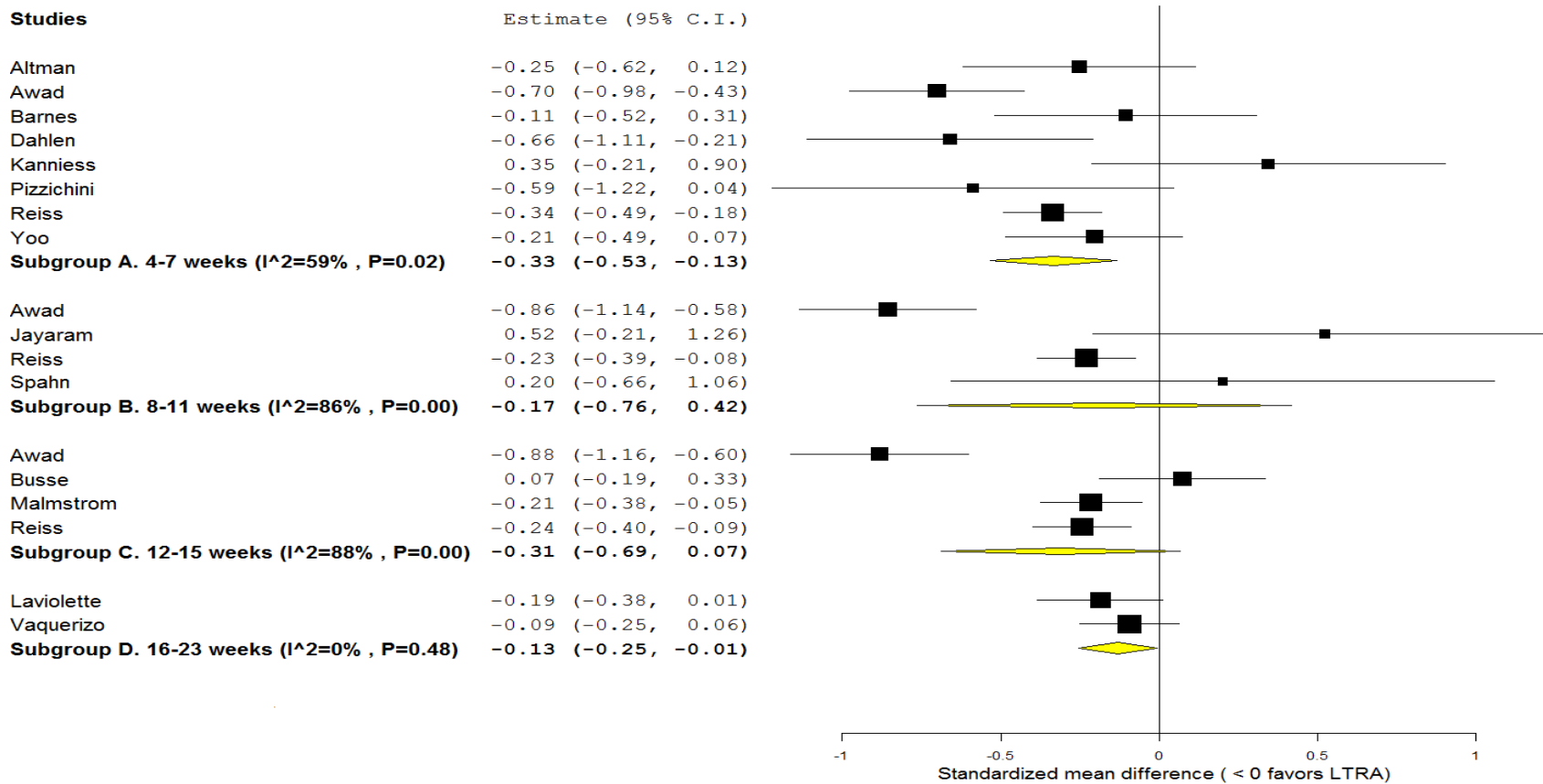


Figure 22. Summary forest plot of standardized mean difference by subgroups of ICS use for daytime symptoms

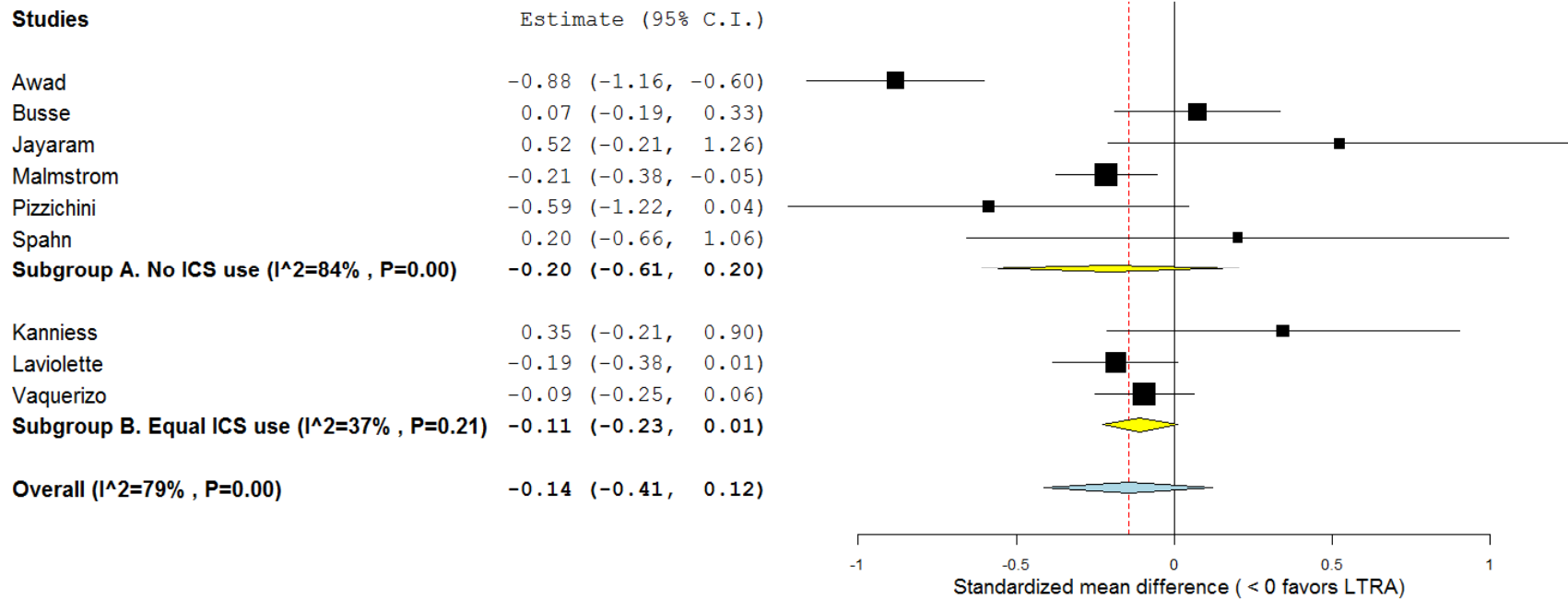


Figure 23. Summary forest plot of mean difference for SABA use

| Studies | Estimate (95% C.I.) |
|--|-----------------------------|
| Altman 1998 | -1.10 (-1.76, -0.44) |
| Barnes 1997 | -0.33 (-1.28, 0.62) |
| Busse 2001 | -0.60 (-1.30, 0.10) |
| Dahlen 2002 | -0.97 (-1.82, -0.12) |
| Fish 1997 | -0.77 (-1.08, -0.46) |
| Jayaram 2005 | -0.20 (-1.84, 1.44) |
| Kanniess 2002 | -0.06 (-0.68, 0.56) |
| Nathan 1998 | -0.64 (-0.91, -0.37) |
| Reid (A) 2008 | 0.80 (-0.59, 2.19) |
| Reid (B) 2008 | -0.80 (-2.91, 1.31) |
| Spahn 2006 | -0.63 (-1.47, 0.21) |
| Spector 1994 | -1.00 (-2.00, -0.00) |
| Overall (I²=7% , P=0.37) | -0.65 (-0.82, -0.49) |

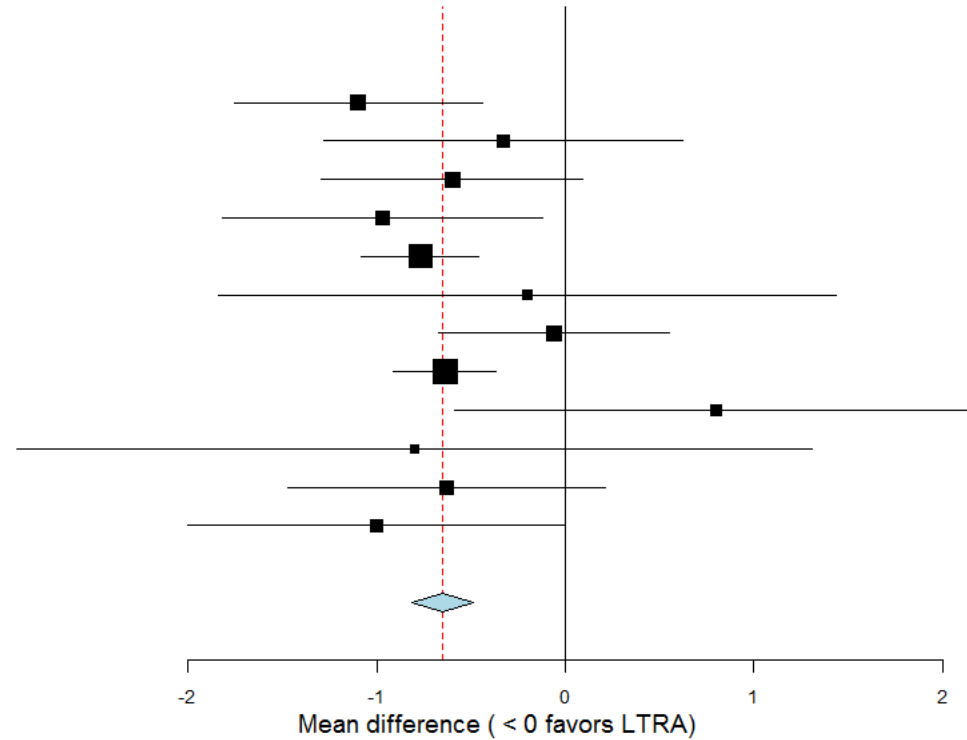


Figure 24. Meta-regression plot of mean difference for SABA by treatment duration

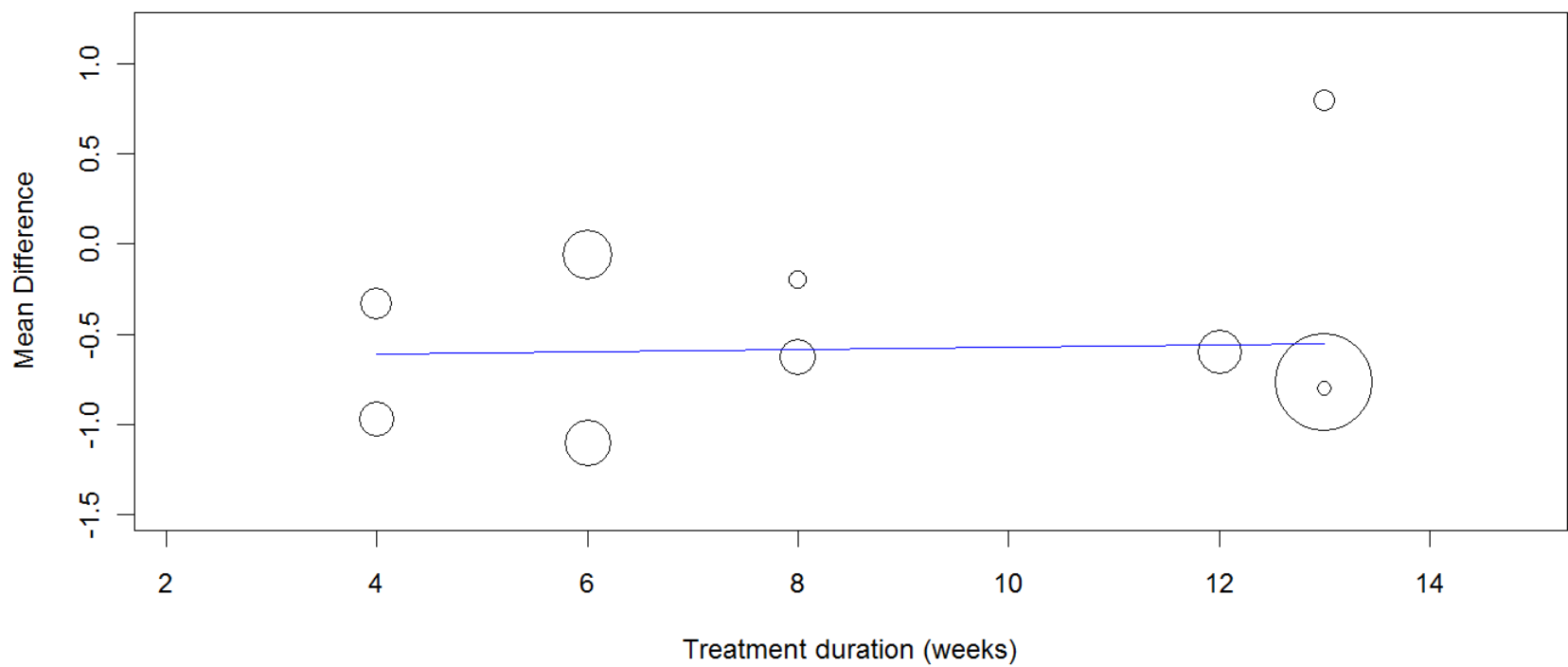


Figure 25. Summary forest plot of mean difference by subgroups of treatment duration for SABA use

| Studies | Estimate (95% C.I.) |
|---|-----------------------------|
| Altman | -1.10 (-1.76, -0.44) |
| Dahlen | -0.97 (-1.82, -0.12) |
| Kanniess | -0.06 (-0.68, 0.56) |
| Jayaram | -0.20 (-1.84, 1.44) |
| Spahn | -0.63 (-1.47, 0.21) |
| Subgroup A. Montelukast (I²=35% , P=0.19) | -0.63 (-1.10, -0.16) |
| Spector | -1.00 (-2.00, -0.00) |
| Busse | -0.60 (-1.30, 0.10) |
| Fish | -0.77 (-1.08, -0.46) |
| Nathan | -0.64 (-0.91, -0.37) |
| Reid (A) | 0.80 (-0.59, 2.19) |
| Reid (B) | -0.80 (-2.91, 1.31) |
| Subgroup B. Zafirlukast (I²=5% , P=0.39) | -0.67 (-0.86, -0.48) |
| Barnes | -0.33 (-1.28, 0.62) |
| Subgroup C. Pranlukast (I²=NA , P=NA) | -0.33 (-1.28, 0.62) |
| Overall (I²=7% , P=0.37) | -0.65 (-0.82, -0.49) |

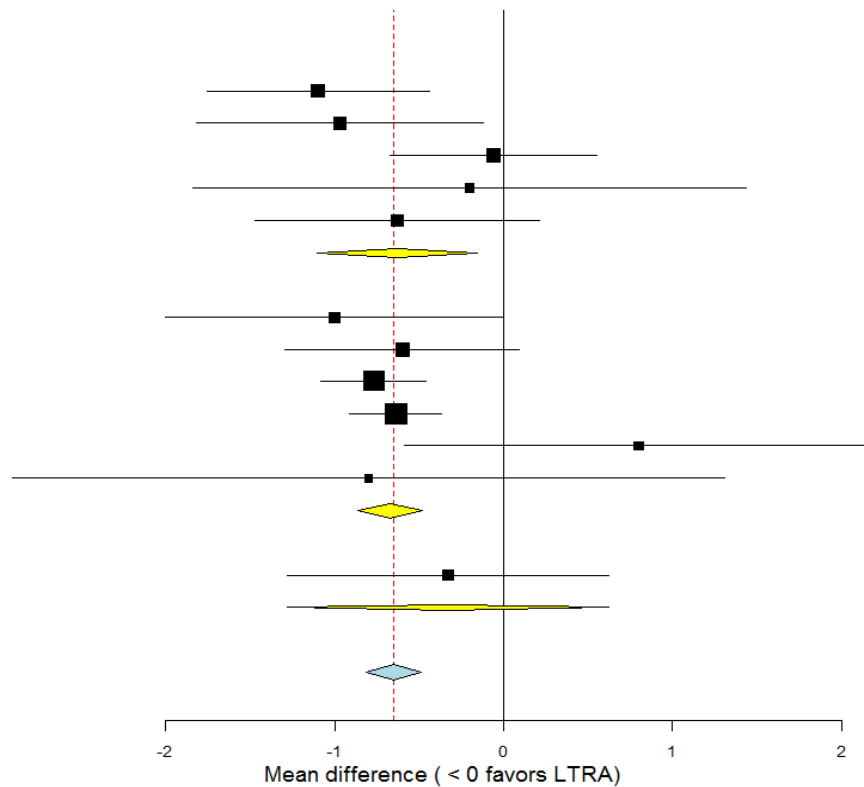


Figure 26. Summary forest plot of mean difference by subgroups of LTRA for SABA use

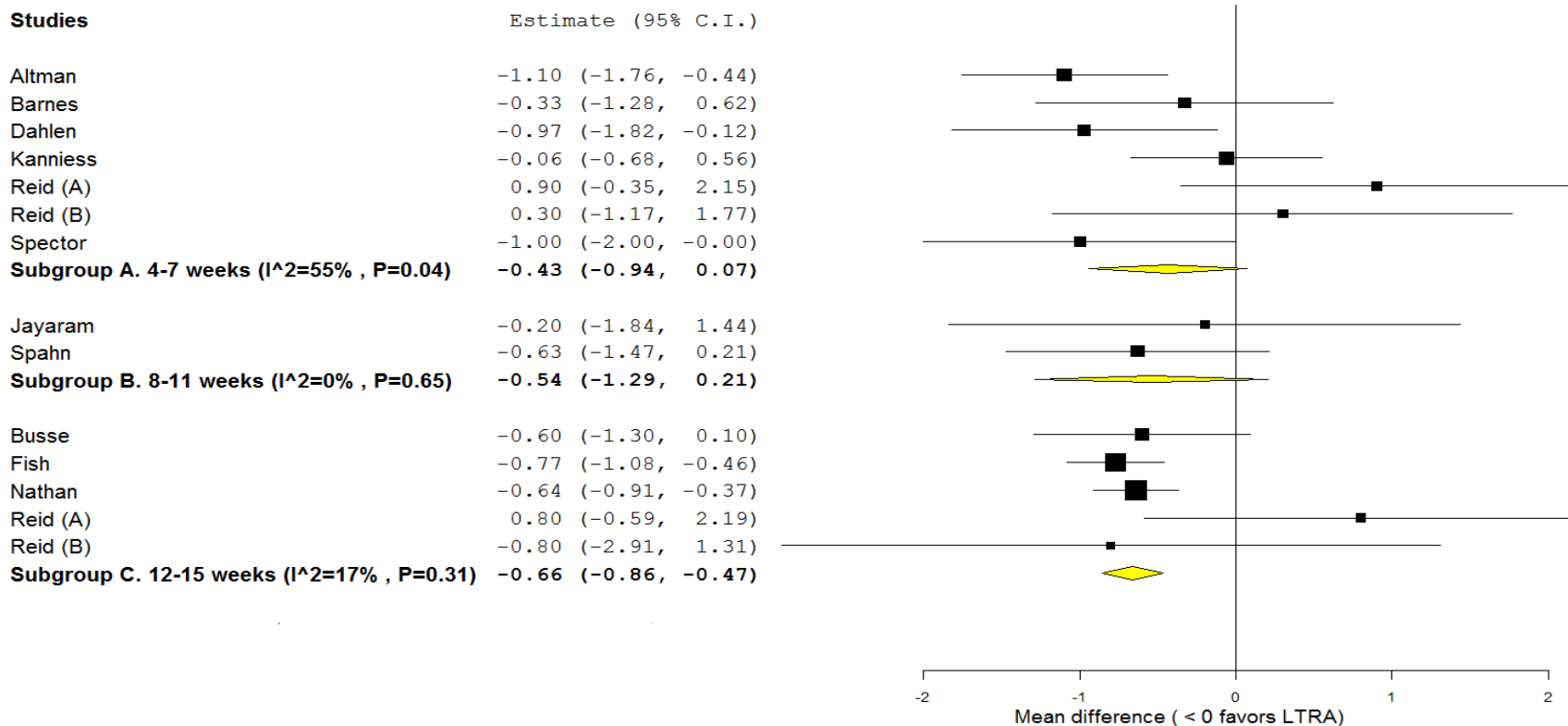


Figure 27. Summary forest plot of mean difference in percent change from baseline for SABA use

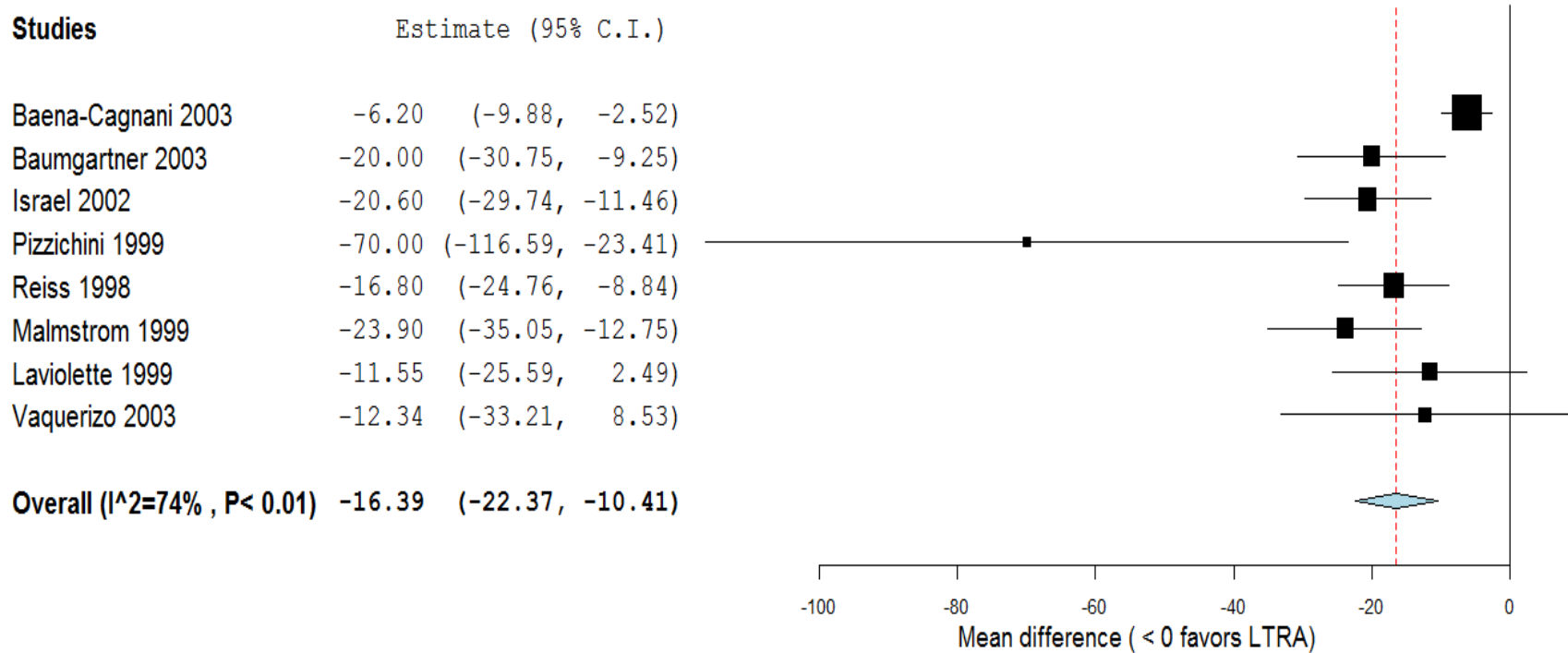


Figure 28. Summary forest plot of mean difference in % change from baseline by subgroups of treatment duration for SABA

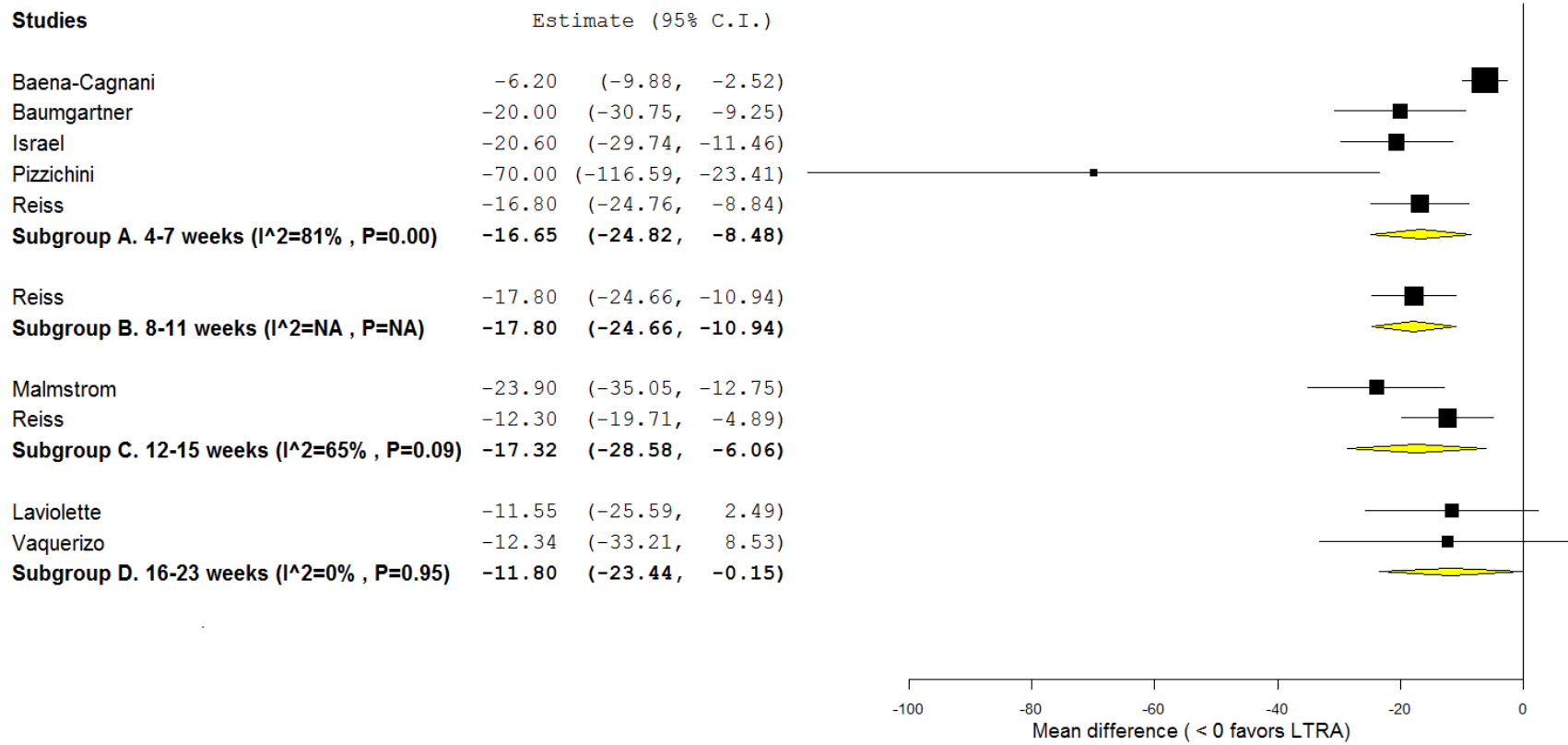


Figure 29. Summary forest plot for mean difference for nocturnal awakenings

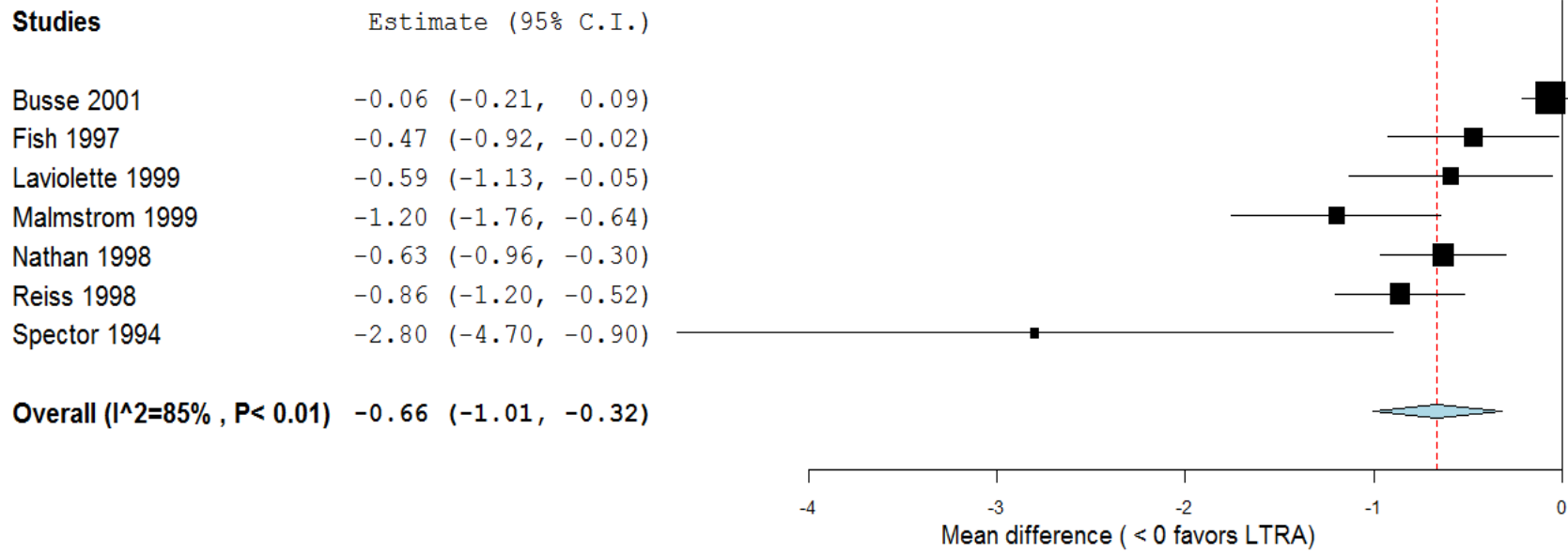
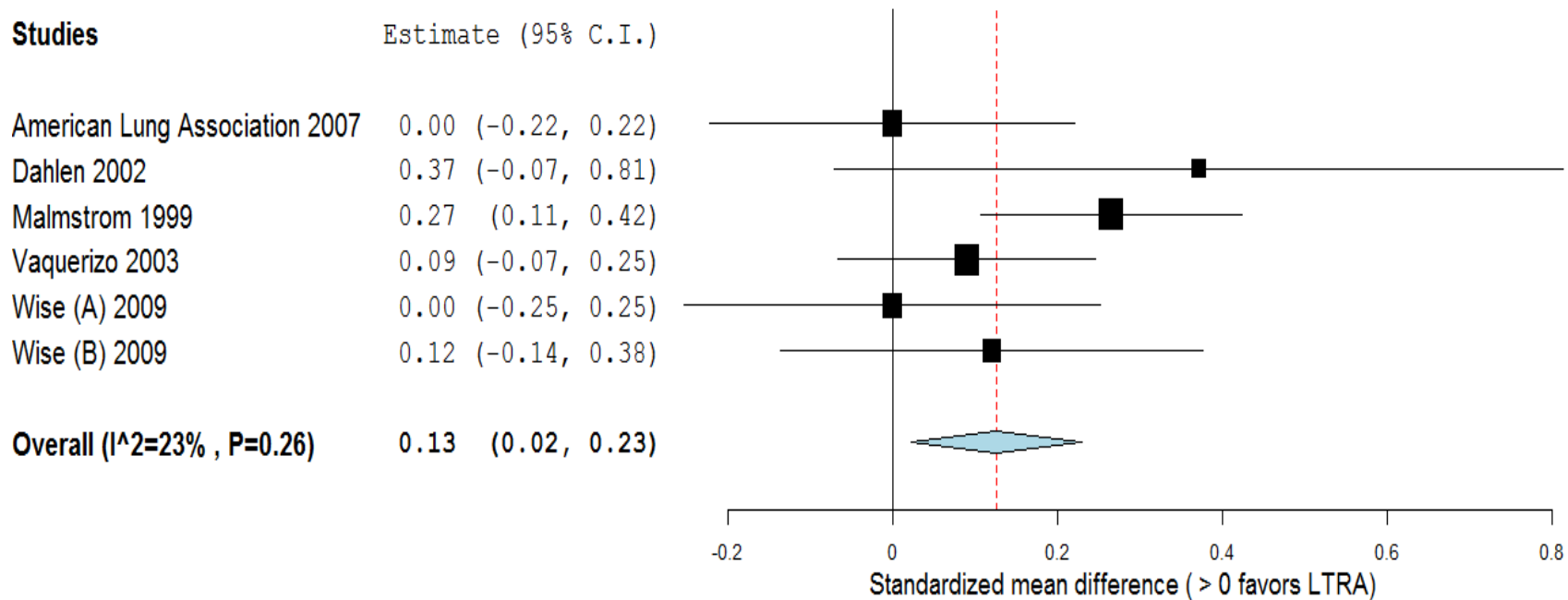


Figure 30. Summary forest plot of standardized mean difference for ASQL



Adverse events and tolerability

The proportions of patients with adverse events were generally similar between intervention and comparator groups (Table 9). Overall, no serious adverse events were reported. Five trials reported no adverse events. Withdrawals from adverse events or worsening asthma were generally similar across our comparator groups (Table 10). The definitions of exacerbations that led to withdrawals are also presented in this table.

Table 9. Adverse events reported in included trials

| Adverse event / Source | Treatment arms | Number of events | Number of patients with ≥ 1 event | Total number of patients | Reported p-value | Comments in original studies |
|-------------------------|-----------------|------------------|--|--------------------------|-------------------------|---|
| Abdominal pain | | | | | | |
| Barnes 1997 | Pranlukast | | 2 | 46 | | |
| | Placebo | | 0 | 44 | | |
| Nathan 1998 | Zafirlukast | | 1 | 125 | | Led to withdrawal |
| | Placebo | | 0 | 132 | | |
| Vaquerizo | Montelukast/ICS | | 8 | 317 | 0.6 | |
| | Placebo/ICS | | 6 | 308 | | |
| Yoo 2001 | Pranlukast | 0 | | 98 | | |
| | Placebo | 1 | | 99 | | |
| Back pain | | | | | | |
| Fish 1997 | Zafirlukast | | 15 | 514 | | |
| | Placebo | | | | | |
| Nathan 2005 | Montelukast | | 8 | 248 | | |
| | Placebo | | | | | |
| Bronchitis | | | | | | |
| Barnes 1997 | Pranlukast | | 1 | 46 | | |
| | Placebo | | 2 | 44 | | |
| Laviolette 1999 | Montelukast/ICS | | 5 | 193 | | |
| | Placebo/ICS | | 4 | 200 | | |
| | Montelukast | | 7 | 201 | | |
| | Placebo | | 4 | 48 | | |
| Vaquerizo 2003 | Montelukast/ICS | | 5 | 317 | 0.2 | |
| | Placebo/ICS | | 3 | 308 | | |
| Chest congestion | | | | | | |
| Busse 2001 | Zafirlukast | | 6 | 111 | | |
| | Placebo | | 0 | 114 | | |
| Clinical AE | | | | | | |
| Altman 1998 | Montelukast | | | 57 | NS | Headache, URTI most commonly observed |
| | Placebo | | | 58 | | |
| Israel 2002 | Montelukast | | | 339 | | Similar frequency, URTI, headache, sinusitis |
| | Placebo | | | 111 | | |
| Löfdahl 1999 | Montelukast | | | | | No significant differences between groups |
| | Placebo | | | | | |
| Minoguchi 2002 | Montelukast | 26 | 13 | 27 | NS | Respiratory events most commonly observed |
| | Placebo | 30 | 15 | 28 | | |
| Nathan 1999 | Zafirlukast | | 64 | 96 | p <0.05 for pharyngitis | Pharyngitis and headache were common |
| | Placebo | | 72 | 95 | | |
| Pizzichini 1999 | Montelukast | | | 19 | | few AE and similar frequencies |
| | Placebo | | | 21 | | |
| Tohda 2002 | Montelukast | | 6 | 89 | 0.6 | headache, stomach ache, heartburn, diarrhea, constipation |
| | Placebo | | 6 | 92 | | |
| Cough | | | | | | |
| Barnes 1997 | Pranlukast | | 0 | 46 | | |
| | Placebo | | 1 | 44 | | |
| Fish 1997 | Zafirlukast | | 16 | 514 | | |
| | Placebo | | 11 | 248 | | |

Table 9. (continued)

| Adverse event / Source | Treatment arms | Number of events | Number of patients with ≥ 1 event | Total number of patients | Reported p-value | Comments in original studies |
|----------------------------|-----------------|------------------|--|--------------------------|------------------|---|
| Laviolette 1999 | Montelukast/ICS | | 8 | 193 | | |
| | Placebo/ICS | | 42 | 200 | | |
| | Montelukast | | 11 | 201 | | |
| | Placebo | | 3 | 48 | | |
| Diarrhea | | | | | | |
| Barnes 1997 | Pranlukast | | 2 | 46 | | |
| | Placebo | | 0 | 44 | | |
| Fish 1997 | Zafirlukast | | 14 | 514 | | |
| | Placebo | | 11 | 248 | | |
| Dyspepsia | | | | | | |
| Nathan 2005 | Montelukast | | 11 | 282 | | |
| | Placebo | | 6 | 290 | | |
| Yoo 2001 | Pranlukast | 1 | | 98 | | |
| | Placebo | 8 | | 99 | | |
| Dyspnea | | | | | | |
| Barnes 1997 | Pranlukast | | 2 | 46 | | |
| | Placebo | | 0 | 44 | | |
| Elevated ALT or AST | | | | | | |
| Baumgartner 2003 | Montelukast | | 1 | 313 | | |
| | Placebo | | 0 | 103 | | |
| Dahlen 2002 | Montelukast | | 5 | 40 | | |
| | Placebo | | 7 | 40 | | |
| Fish 1997 | Zafirlukast | | 17 | 514 | NS | |
| | Placebo | | 10 | 248 | | |
| Leff 1998 | Montelukast | | 1 | 54 | | More than 3 times the upper limit of normal |
| | Placebo | | 1 | 56 | | |
| Malmstrom 1999 | Montelukast | | 85 | 382 | | 4 patients in each arm had levels more than 3 times the upper limit of normal |
| | Placebo | | 56 | 254 | | |
| Nathan 1998 | Zafirlukast | | 2 | 231 | | |
| | Placebo | | 8 | 223 | | |
| Spector 1994 | Zafirlukast | | 3 | 67 | | |
| | Placebo | | 3 | 66 | | |
| Tohda 2002 | Montelukast | | 10 | 89 | | |
| | Placebo | | 8 | 92 | | |
| Yoo 2001 | Pranlukast | 1 | | 98 | | |
| | Placebo | 2 | | 99 | | |
| Epistaxis | | | | | | |
| Nathan 2005 | Montelukast | | 6 | 282 | | |
| | Placebo | | 12 | 290 | | |
| Gastritis | | | | | | |
| Awad 2002 | Zafirlukast | 6 | | 116 | | |
| | Placebo | 2 | | 99 | | |
| Spector 1994 | Zafirlukast | | 1 | 67 | | |
| | Placebo | | 4 | 66 | | |
| Headache | | | | | | |
| Baena-Cagnani 2003 | Montelukast | | 11 | 311 | | |
| | Placebo | | 11 | 302 | | |
| Barnes 1997 | Pranlukast | | 1 | 42 | | |
| | Placebo | | 2 | 43 | | |
| Baumgartner 2003 | Montelukast | | 31 | 313 | | |
| | Placebo | | 18 | 103 | | |

Table 9. (continued)

| Adverse event / Source | Treatment arms | Number of events | Number of patients with ≥ 1 event | Total number of patients | Reported p-value | Comments in original studies |
|------------------------|-----------------|------------------|--|--------------------------|------------------|------------------------------------|
| Busse 2001 | Zafirlukast | | 2 | 111 | | |
| | Placebo | | 3 | 114 | | |
| Dahlen 2002 | Montelukast | | 0 | 40 | | Severe that led to withdrawal |
| | Placebo | | 1 | 40 | | |
| Fish 1997 | Zafirlukast | | 71 | 514 | | |
| | Placebo | | 28 | 248 | | |
| Laviolette 1999 | Montelukast/ICS | | 50 | 193 | | |
| | Placebo/ICS | | 42 | 200 | | |
| | Montelukast | | 52 | 201 | | |
| | Placebo | | 3 | 48 | | |
| Leff 1998 | Montelukast | | 11 | 54 | | |
| | Placebo | | 16 | 56 | | |
| Malmstrom 1999 | Montelukast | | 68 | 387 | | |
| | Placebo | | 40 | 257 | | |
| Nathan 2005 | Montelukast | | 40 | 282 | | |
| | Placebo | | 38 | 290 | | |
| Reiss 1998 | Montelukast | | 73 | 408 | | |
| | Placebo | | 57 | 273 | | |
| Spector 1994 | Zafirlukast | | 5 | 68 | | |
| | Placebo | | 8 | 70 | | |
| Vaquerizo 2003 | Montelukast | | 34 | 317 | 0.6 | |
| | Placebo | | 29 | 308 | | |
| Wise (A) 2009 | Montelukast | | 44 | 120 | | |
| | Placebo | | 34 | 121 | | |
| Wise (B) | Montelukast | | 35 | 119 | | |
| | Placebo | | 23 | 120 | | |
| Yoo 2001 | Pranlukast | 1 | | 98 | | |
| | Placebo | 4 | | 99 | | |
| Hepatitis | | | | | | |
| Jayaram 2005 | Montelukast | 1 | | 14 | | Drug induced |
| | Placebo | 0 | | | | |
| Jayaram 2005 | Montelukast | 1 | | 19 | | Drug induced |
| | Placebo | 0 | | 22 | | |
| Hypertonia | | | | | | |
| Fish 1997 | Zafirlukast | | 15 | 514 | | |
| | Placebo | | 8 | 248 | | |
| Nathan 1998 | Zafirlukast | | 1 | 125 | | Led to withdrawal |
| | Placebo | | 0 | 132 | | |
| Influenza | | | | | | |
| Fish 1997 | Zafirlukast | | 17 | 514 | | |
| | Placebo | | 11 | 248 | | |
| Laviolette 1999 | Montelukast/ICS | | 11 | 193 | | |
| | Placebo/ICS | | 11 | 200 | | |
| | Montelukast | | 15 | 201 | | |
| | Placebo | | 3 | 48 | | |
| Malmstrom 1999 | Montelukast | | 25 | 387 | | |
| | Placebo | | 10 | 257 | | |
| Vaquerizo 2003 | Montelukast | | 38 | 317 | 0.7 | |
| | Placebo | | 34 | 308 | | |
| Laboratory AE | | | | | | |
| Altman 1998 | Montelukast | | | | | More frequent in the placebo group |
| | Placebo | | | | | |
| Israel 2002 | Montelukast | | 13 | 339 | NS | |
| | Placebo | | 5 | 111 | | |

Table 9. (continued)

| Adverse event / Source | Treatment arms | Number of events (%) | Number of patients with ≥ 1 event (%) | Total number of patients | Reported p-value | Comments in original studies |
|---------------------------------------|-----------------|----------------------|--|--------------------------|------------------|--|
| Löfdahl 1999 | Montelukast | | | 113 | | No significant differences between groups |
| | Placebo | | | 113 | | |
| Minoguchi 2002 | Montelukast | 5 | 4 | 27 | NS | |
| | Placebo | 6 | 5 | 28 | | |
| Nathan 1998 | Montelukast | | 5 | 125 | | Increased bilirubin or alkaline phosphatase |
| | Placebo | | 1 | 132 | | |
| Pizzichini 1999 | Montelukast | | | | NS | Infrequent and similar frequencies leukocytosis, increased levels of liver enzymes, glycosuria |
| | Placebo | | | | | |
| Tohda 2002 | Montelukast | | 12 | 89 | NS | |
| | Placebo | | 7 | 92 | | |
| Yoo 2001 | Pranlukast | 2 | | 98 | | Increased bilirubin |
| | Placebo | 1 | | 99 | | |
| Myalgia | | | | | | |
| Fish 1997 | Zafirlukast | | 19 | 514 | | |
| | Placebo | | 9 | 248 | | |
| Nausea | | | | | | |
| American Lung Association 2007 | Montelukast | | 44 | 164 | | |
| | Placebo | | 52 | 164 | | |
| Awad 2002 | Zafirlukast | 5 | | 116 | | |
| | Placebo | 1 | | 99 | | |
| Laviolette 1999 | Montelukast/ICS | | 5 | 193 | | |
| | Placebo/ICS | | 11 | 200 | | |
| | Montelukast | | 12 | 201 | | |
| | Placebo | | 0 | 48 | | |
| Yoo 2001 | Pranlukast | 1 | | 98 | | |
| | Placebo | 1 | | 99 | | |
| Nervousness | | | | | | |
| American Lung Association 2007 | Montelukast | | 62 | 164 | | |
| | Placebo | | 57 | 164 | | |
| Oropharyngeal candidiasis | | | | | | |
| Busse 2001 | Zafirlukast | | 0 | 111 | | |
| | Placebo | | 2 | 114 | | |
| Pharyngitis | | | | | | |
| Awad 2002 | Zafirlukast | 2 | | 116 | | |
| | Placebo | 2 | | 99 | | |
| Barnes 1997 | Pranlukast | | 1 | 46 | | |
| | Placebo | | 0 | 44 | | |
| Fish 1997 | Zafirlukast | | 127 | 514 | | |
| | Placebo | | 53 | 248 | | |
| Laviolette 1999 | Montelukast/ICS | | 10 | 193 | | |
| | Placebo/ICS | | 16 | 200 | | |
| | Montelukast | | 12 | 201 | | |
| | Placebo | | 2 | 48 | | |
| Malmstrom 1999 | Montelukast | | 25 | 387 | | |
| | Placebo | | 11 | 257 | | |
| Nathan 1998 | Zafirlukast | | 43 | 231 | | |
| | Placebo | | 42 | 223 | | |
| Reiss 1998 | Montelukast | | 22 | 408 | | |
| | Placebo | | 29 | 273 | | |
| Spector 1994 | Zafirlukast | | 14 | 68 | | |
| | Placebo | | 16 | 70 | | |

Table 9. (continued)

| Adverse event / Source | Treatment arms | Number of events (%) | Number of patients with ≥ 1 event (%) | Total number of patients | Reported p-value | Comments in original studies |
|--|-----------------|----------------------|--|--------------------------|------------------|------------------------------|
| Vaquerizo 2003 | Montelukast | | 5 | 317 | 0.8 | |
| | Placebo | | 4 | 308 | | |
| Rash/Itching | | | | | | |
| Barnes 1997 | Pranlukast | | 1 | 46 | | |
| | Placebo | | 1 | 44 | | |
| Fish 1997 | Zafirlukast | | 16 | 514 | | |
| | Placebo | | 9 | 248 | | |
| Laviolette 1999 | Montelukast/ICS | | 1 | 193 | | |
| | Placebo/ICS | | 3 | 200 | | |
| | Montelukast | | 7 | 201 | | |
| | Placebo | | 3 | 48 | | |
| Yoo 2001 | Pranlukast | 7 | | 98 | | |
| | Placebo | 4 | | 99 | | |
| Respiratory disorder | | | | | | |
| Barnes 1997 | Pranlukast | | 3 | 46 | | |
| | Placebo | | 1 | 44 | | |
| Rhinitis | | | | | | |
| Awad 2002 | Zafirlukast | 1 | | 116 | | |
| | Placebo | 6 | | 99 | | |
| Fish 1997 | Zafirlukast | | 16 | 514 | | |
| | Placebo | | 8 | 248 | | |
| Nathan 1998 | Zafirlukast | | 5 | 231 | | |
| | Placebo | | 8 | 223 | | |
| Spector 1994 | Zafirlukast | | 5 | 68 | | |
| | Placebo | | 1 | 70 | | |
| Vaquerizo 2003 | Montelukast | | 5 | 317 | 0.7 | |
| | Placebo | | 6 | 308 | | |
| Sinusitis | | | | | | |
| Busse 2001 | Zafirlukast | | 4 | 111 | | |
| | Placebo | | 4 | 114 | | |
| Fish 1997 | Zafirlukast | | 18 | 514 | | |
| | Placebo | | 12 | 248 | | |
| Laviolette 1999 | Montelukast/ICS | | 8 | 193 | | |
| | Placebo/ICS | | 9 | 200 | | |
| | Montelukast | | 12 | 201 | | |
| | Placebo | | 2 | 48 | | |
| Nathan 1998 | Zafirlukast | | 8 | 231 | | |
| | Placebo | | 13 | 223 | | |
| Reiss 1998 | Montelukast | | 31 | 408 | | |
| | Placebo | | 22 | 273 | | |
| Sore throat | | | | | | |
| Busse 2001 | Zafirlukast | | 3 | 111 | | |
| | Placebo | | 3 | 114 | | |
| Nathan 2005 | Montelukast | | 11 | 282 | | |
| | Placebo | | 9 | 290 | | |
| Upper respiratory tract infection | | | | | | |
| Baumgartner 2003 | Montelukast | | 22 | 313 | | |
| | Placebo | | 7 | 103 | | |
| Laviolette 1999 | Montelukast/ICS | | 70 | 193 | | |
| | Placebo/ICS | | 79 | 200 | | |
| | Montelukast | | 72 | 201 | | |
| | Placebo | | 20 | 48 | | |
| Leff 1998 | Montelukast | | 12 | 54 | | |
| | Placebo | | 16 | 56 | | |

Table 9. (continued)

| Adverse event / Source | Treatment arms | Number of events | Number of patients with ≥ 1 event (%) | Total number of patients | Reported p-value | Comments in original studies |
|--------------------------------|----------------|------------------|--|--------------------------|------------------|------------------------------|
| Malmstrom 1999 | Montelukast | | 48 | 387 | | |
| | Placebo | | 28 | 357 | | |
| Reiss 1999 | Montelukast | | 129 | 408 | | |
| | Placebo | | 96 | 273 | | |
| Vaquerizo 2003 | Montelukast | | 17 | 317 | 0.5 | |
| | Placebo | | 21 | 308 | | |
| Urinary Tract infection | | | | | | |
| Vaquerizo 2003 | Montelukast | | 6 | 317 | 0.7 | |
| | Placebo | | 7 | 308 | | |
| Vomiting | | | | | | |
| Awad 2002 | Zafirlukast | 1 | | 116 | | |
| | Placebo | 0 | | 99 | | |

Abbreviations: AE: Adverse events; ALT: Alanine Aminotransferase; AST: Aspartate Aminotransferase; NS: Not statistically significant; URTI: Upper Respiratory Tract Infection.

Table 10. Withdrawals in included trials

| Source | Treatment arms | Number of withdrawals due to adverse events | Number of withdrawals due to worsening asthma / exacerbation | Total number of patients | Comments about exacerbations |
|---------------------------------------|-----------------|---|--|--------------------------|---|
| American Lung Association 2007 | Montelukast | 7 | | 164 | |
| | Placebo | 5 | | 164 | |
| Barnes 1997 | Pranlukast | | 2 | 46 | No definition of an exacerbation reported |
| | Placebo | | 1 | 44 | |
| Baumgartner 2003 | Montelukast | 1 | | 313 | |
| | Placebo | 3 | | 103 | |
| Busse 2001 | Zafirlukast | 1 | | 111 | |
| | Placebo | 1 | | 114 | |
| Dahlén 2002 | Montelukast | 0 | | 40 | |
| | Placebo | 1 | | 40 | |
| Fish 1997 | Zafirlukast | 12 | 16 | 514 | No definition of an exacerbation reported |
| | Placebo | 7 | 16 | 248 | |
| Huang 2003 | Zafirlukast | | 3 | 20 | Exacerbation that required emergency room visit |
| | Placebo | | 4 | 18 | |
| Kanniess 2002 | Montelukast | | 3 | 26 | An exacerbation was defined as a decrease in PEF of >50% compared to values at entry, or an increase in daytime symptoms of three or more or in night-time symptoms of two or more on >3 consecutive days |
| | Placebo | | 2 | 24 | |
| Laviolette 1999 | Montelukast/ICS | | 2 | 193 | |
| | Placebo/ICS | | 8 | 200 | |
| | Montelukast | | 23 | 201 | |
| | Placebo | | 7 | 48 | |
| Leff 1998 | Montelukast | 2 | | 54 | |
| | Placebo | 4 | | 56 | |
| Löfdahl 1999 | Montelukast | 0 | 4 | 112 | Exacerbation that required oral corticosteroids |
| | Placebo | 9 | 0 | 113 | |
| Malmstrom 1999 | Montelukast | 8 | | | |
| | Placebo | 11 | | | |
| Minoguchi 2002 | Montelukast | | 1 | 27 | During wash-out period |
| | Placebo | | 2 | 28 | |
| Nakamura 1998 | Pranlukast | 0 | | 11 | |
| | Placebo | 2 | | 10 | |
| Nathan 1998 | Zafirlukast | 3 | 2 | 231 | |
| | Placebo | 0 | 6 | 223 | |
| Nathan 1999 | Zafirlukast | | 3 | 96 | No definition of an exacerbation reported |
| | Placebo | | 8 | 95 | |
| Nathan 2005 | Montelukast | | 1 | 282 | Exacerbation defined as any event that required treatment with asthma medications beyond study medications |
| | Placebo | | 2 | 290 | |
| Pizzichini 1999 | Montelukast | 1 | 1 | 19 | |
| | Placebo | 1 | 0 | 21 | |
| Reiss 1998 | Montelukast | | 6 | 408 | |
| | Placebo | | 10 | 273 | |
| Spahn 2006 | Montelukast | | 1 | 11 | Required rescue prednisone |
| | Placebo | | 1 | 10 | |
| Spector 1994 | Zafirlukast | | 0 | 70 | Treated with no specific treatment protocol |
| | Placebo | | 8 | 70 | |

Table 10. (continued)

| Source | Treatment arms | Number of withdrawals due to adverse events | Number of withdrawals due to worsening asthma / exacerbation | Total number of patients | Comments about exacerbations |
|--------------------------|----------------|---|--|--------------------------|---|
| Stelmach (A) 2007 | Montelukast | | 0 | 29 | |
| | Placebo | | 0 | 29 | |
| Stelmach (B) | Montelukast | | 0 | 29 | No definition of an exacerbation reported |
| | Placebo | | 2 | 29 | |
| Strunk 2008 | Montelukast | | 0 | 19 | Exacerbations required oral corticosteroids |
| | Placebo | | 3 | 19 | |
| Vaquerizo 2003 | Montelukast | 3 | 3 | 317 | |
| | Placebo | 5 | 3 | 308 | |
| Wise (A) 2009 | Montelukast | 2 | | 120 | |
| | Placebo | 2 | | 121 | |
| Wise (B) | Montelukast | 2 | | 120 | |
| | Placebo | 2 | | 121 | |

Discussion

Administration of a LTRA to adults and adolescents with asthma significantly reduced the risk of an exacerbation and improved both lung function and patient-reported outcomes compared to placebo. This effect was consistent across all types of LTRAs. The effect on the risk of exacerbations was more evident in studies of shorter duration. The incidence of adverse events and withdrawals due to adverse events and worsening asthma was similar for LTRAs and placebo reflecting a favorable safety and tolerability profile for LTRAs.

Several systematic reviews have examined the state of evidence regarding the use of LTRAs in adults and adolescents with asthma, but only a few have included RCTs that have compared a LTRA with placebo. Joos et al.⁶⁶ included RCTs of at least 12 weeks duration that examined the benefits and harms of montelukast as add-on therapy to ICS compared to ICS with or without placebo and concluded that the addition of montelukast to ICS improved control of mild to moderate asthma compared to ICS monotherapy; no meta-analytic technique was employed in this study, however, due to the inclusion of a limited number of RCTs. Ducharme et al.⁶⁷ included RCTs of at least 4 weeks duration that compared LTRAs with placebo as add-on to ICS, but only two out of six included RCTs reported use of usual licensed doses. Currie et al.⁶⁸ examined the bronchoprotective effects of LTRAs compared to placebo after administration of bronchial stimuli. These provocative challenges, though, are mainly used in order to establish the diagnosis of asthma. In contrast, our systematic review was more expansive and more applicable to current clinical practice in a number of ways. Our study included outcome measures that correspond to the components proposed by international guidelines to periodically assess

and monitor asthma control in patients with an established diagnosis of asthma.⁶⁹ More specifically, we included RCTs of at least 4 weeks duration because the level of asthma control is assessed over a 4 week period at the minimum.⁶ We also excluded RCTs where LTRAs were not administered on a daily basis because we intended to examine their effect as long-term controller medications.⁶ Therefore, RCTs that only assessed the pharmacodynamic profile of single doses of LTRAs after provocative challenges or exercise were excluded. In a systematic review of LTRA safety data that included both RCTs and their extension studies, Storms et al.⁷⁰ concluded that there was no significant difference in the incidence of adverse events between patients who were treated with montelukast and those who received placebo. Although our systematic review included only RCTs with a relatively short length of follow-up, the reported adverse events and their relative frequencies were similar to those reported in the meta-analysis by Storms et al.

Despite the broader scope and improved generalizability of our study to previous reviews, our systematic review had several limitations both at the individual-study level and the systematic review level that need to be kept in mind in interpreting our study results. Different definitions of asthma exacerbation were used in the trials reviewed and some studies did not explicitly define this important clinical endpoint. Therefore, the severity of the exacerbations could not be fully assessed across studies. Importantly, in 12 studies only withdrawals due to an exacerbation were reported. Taking the conservative approach of meta-analyzing asthma exacerbations only from studies that clearly defined and reported the outcome, we assumed that no events occurred in studies that did not mention any exacerbation and we did not quantitatively summarize those studies which

reported only withdrawals because the number of patients at risk could not be determined. In addition, the definitions of asthma severity in the individual trials were not consistent and it was not possible to assess the impact of baseline asthma severity on summary treatment effects. Our conclusions about the magnitude of the treatment effect are limited due to the small number of RCTs included in each analysis compared to the total number of RCTs included in the systematic review. Part of this difference arises because studies used different analytic scales for the same outcome measure, which prevented combination of all available data. For instance, change from baseline in FEV₁ was provided as an absolute number in some trials and as a percentage in others. Another set of studies reported a non-significant difference, without providing the actual numbers, for clinical outcomes. These studies were generally either not primarily designed to assess outcome measures relevant to clinical practice or had a relatively small sample size. In either case, this leads to outcome reporting bias and so the summary effect sizes from the meta-analyses may be overstated.

The large amount of between-study statistical heterogeneity found for most outcomes could be sometimes be partly explained by subgroup analyses and meta-regression, but such analyses can only be hypothesis generating by the retrospective nature of meta-analysis. For other outcomes, however, the observed statistical heterogeneity remained largely unexplained. Potential association between allergic rhinitis and the magnitude of the summary treatment effect remains unclear due to insufficient reporting in the individual studies. Similarly, only two trials reported inclusion of patients with aspirin-induced asthma. Another limitation of the study is the inclusion of only peer-reviewed and English-language publications. Finally, we cannot exclude publication bias.

To our knowledge, this is the first systematic review to investigate the use of all marketed LTRAs in usual licensed doses as asthma controller medications compared with placebo. Our findings suggest that LTRAs might be an efficacious and safe alternative treatment, in adult and adolescent patients who cannot or prefer not to take ICS. However, which patients are more likely to respond to LTRAs administered as monotherapy remains unclear. Asthma is a complex disease with various clinical, inflammatory, and trigger-related “phenotypes” that may overlap.⁷¹ It is hypothesized that proper identification of these phenotypes would lead to better management of the disease. Moreover, since not all patients respond well to ICS, the need for alternative treatments that would benefit specific subpopulations increases. Therefore, professional organizations or expert panels should recommend standardized study-protocols, definitions of phenotypes, and outcome measures (e.g., asthma control test, asthma control questionnaire)⁷²⁻⁷³ for the purpose of research and encourage future researchers to implement these standards.

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